

# Asymptotic Efficiency in High-Dimensional Covariance Estimation

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Lübbenau, March 2018

Let  $X, X_1, \dots, X_n$  be i.i.d. Gaussian vectors with values in  $\mathbb{R}^d$ , with  $\mathbb{E}X = 0$  and with covariance operator  $\Sigma = \mathbb{E}(X \otimes X) \in \mathcal{C}_+^d$ .

- Given a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$  and a linear operator  $B : \mathbb{R}^d \mapsto \mathbb{R}^d$  with  $\|B\|_1 \leq 1$ , estimate  $\langle f(\Sigma), B \rangle$  based on  $X_1, \dots, X_n$ .
- More precisely, we are interested in finding **asymptotically efficient** estimators of  $\langle f(\Sigma), B \rangle$  with  $\sqrt{n}$ -convergence rate in the case when  $d = d_n \rightarrow \infty$ .
- Suppose  $d_n \leq n^\alpha$  for some  $\alpha > 0$ . Is there  $s(\alpha)$  such that for all  $s > s(\alpha)$  and for all functions  $f$  of smoothness  $s$ , asymptotically efficient estimation is possible?

# Some Related Results

- Efficient estimation of smooth functionals in nonparametric models: Levit (1975, 1978), Ibragimov and Khasminskii (1981);
- In particular, in Gaussian shift model: Ibragimov, Nemirovski and Khasminskii (1987), Nemirovski (1990, 2000)
- Girko (1987–): asymptotically normal estimators of a number of special functionals (such as  $\log \det(\Sigma) = \text{tr}(\log \Sigma)$ , Stieltjes transform of spectral function of  $\Sigma : \text{tr}((I + t\Sigma)^{-1})$ ), ... Based on martingale CLT
- Asymptotic normality of log-determinant  $\log \det(\hat{\Sigma})$  has been studied by many authors (see, e.g., Cai, Liang and Zhou (2015) for a recent result)

# Some Related Results

- Asymptotic normality of  $\text{tr}(f(\hat{\Sigma}))$  for a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$  : (linear spectral statistic). Common topic in random matrix theory (both for Wigner and for Wishart matrices): Bai and Silverstein (2004), Lytova and Pastur (2009), Sosoe and Wong (2015)
- Estimation of functionals of covariance matrices under sparsity: Fan, Rigollet and Wang (2015)
- Bernstein–von Mises theorems for functionals of covariance: Gao and Zhou (2016)
- Efficient estimation of linear functionals of principal components: Koltchinskii, Löffler and Nickl (2017)

- **Part 1.**  
Effective Rank and Sample Covariance
- **Part 2.**  
Taylor Expansions of Operator Functions and Normal Approximation of Plug-In Estimators of Smooth Functionals of Covariance
- **Part 3.**  
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Asymptotic Efficiency

# Part 1.

## Effective Rank and Sample Covariance

# Covariance Operator

- $(E, \|\cdot\|)$  a separable Banach space,  $E^*$  its dual space
- $X$  a centered random variable in  $E$ ,  $\mathbb{E}|\langle X, u \rangle|^2 < +\infty$ ,  $u \in E^*$
- The covariance operator:

$$\Sigma u := \mathbb{E}\langle X, u \rangle X, \quad u \in E^*.$$

- $\Sigma : E^* \mapsto E$  a bounded symmetric nonnegatively definite operator.  
If  $\mathbb{E}\|X\|^2 < +\infty$ , then  $\Sigma$  is nuclear

# Sample Covariance Operator

- $X_1, \dots, X_n$  i.i.d. copies of  $X$ .
- The sample (empirical) covariance operator  $\hat{\Sigma} : E^* \mapsto E$ ,

$$\hat{\Sigma}u := n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j, \quad u \in E^*.$$

- **Problems:**
  - What is the size of  $\mathbb{E}\|\hat{\Sigma} - \Sigma\|$ , where  $\|\cdot\|$  is the operator norm?
  - Concentration inequalities for  $\|\hat{\Sigma} - \Sigma\|$  around its expectation or median.



# Subgaussian Random Variables

## Definition

A centered random variable  $X$  in  $E$  will be called *subgaussian* iff, for all  $u \in E^*$ ,

$$\|\langle X, u \rangle\|_{\psi_2} \lesssim \|\langle X, u \rangle\|_{L_2(\mathbb{P})}.$$

**Notations:** Given a convex nondecreasing function  $\psi : \mathbb{R}_+ \mapsto \mathbb{R}_+$ ,  $\psi(0) = 0$ ,  $\eta$  a r.v. on a probability space  $(\Omega, \mathcal{A}, \mathbb{P})$ ,

$$\|\eta\|_{\psi} := \inf \left\{ C > 0 : \mathbb{E} \psi \left( \frac{|\eta|}{C} \right) \leq 1 \right\}.$$

$$\psi_2(u) := e^{u^2} - 1, u \geq 0$$

$$\psi_1(u) := e^u - 1.$$

# A Bound in the Finite Dimensional Case

$E = \mathbb{R}^d$ ,  $d \geq 1$  (Euclidean space)

## Theorem

Suppose that  $X$  is subgaussian. Then, there exists an absolute constant  $C > 0$  such that, for all  $t \geq 1$ , with probability at least  $1 - e^{-t}$

$$\|\hat{\Sigma} - \Sigma\| \leq C\|\Sigma\| \left( \sqrt{\frac{d}{n}} V \frac{d}{n} V \sqrt{\frac{t}{n}} V \frac{t}{n} \right).$$

It implies that

$$\mathbb{E}\|\hat{\Sigma} - \Sigma\| \leq C\|\Sigma\| \left( \sqrt{\frac{d}{n}} V \frac{d}{n} \right).$$

# Sketch of the proof

- $M \subset S^{d-1}$  is a  $1/4$ -net of the unit sphere  $S^{d-1}$ ,  $\text{card}(M) \leq 9^d$
- 

$$\begin{aligned}\|\hat{\Sigma} - \Sigma\| &\lesssim \max_{u, v \in M} |\langle (\hat{\Sigma} - \Sigma)u, v \rangle| \\ &= \max_{u, v \in M} \left| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle \langle X_j, v \rangle - \mathbb{E} \langle X, u \rangle \langle X, v \rangle \right|\end{aligned}$$

- Use the union bound and Bernstein inequality

$$\begin{aligned}\mathbb{P} \left\{ \left| \frac{\xi_1 + \dots + \xi_n}{n} \right| \gtrsim \|\xi\|_{\psi_1} \left( \sqrt{\frac{t + (2 \log 9)d}{n}} \vee \frac{t + (2 \log 9)d}{n} \right) \right\} \\ \leq \exp\{-t - (2 \log 9)d\}\end{aligned}$$

for independent  $\psi_1$  random variables  $\xi_j := \langle X_j, u \rangle \langle X_j, v \rangle$ .

## Definition

Assuming that  $X$  is a centered Gaussian random variable in  $E$  with covariance operator  $\Sigma$ , define

$$\mathbf{r}(\Sigma) := \frac{\mathbb{E}\|X\|^2}{\|\Sigma\|} = \frac{\mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \langle X, u \rangle \langle X, v \rangle}{\sup_{\|u\|, \|v\| \leq 1} \mathbb{E} \langle X, u \rangle \langle X, v \rangle}.$$

- If  $E$  is a Hilbert space,  $\mathbb{E}\|X\|^2 = \text{tr}(\Sigma)$  and  $\mathbf{r}(\Sigma) = \frac{\text{tr}(\Sigma)}{\|\Sigma\|}$ .
- $\mathbf{r}(\Sigma)$  is called “effective rank” (Vershynin (2012)).
- $\mathbf{r}(\Sigma) \leq \text{rank}(\Sigma)$ .
- If  $\dim(\mathbb{H}) = d < +\infty$  and  $\Sigma$  is of *isotropic type*, that is, for some constants  $0 < c_1 \leq c_2 < \infty$ ,  $c_1 I_d \preceq \Sigma \preceq c_2 I_d$ , then  $\mathbf{r}(\Sigma) \asymp d$ .

# Bounds in Terms of Effective Rank

- **Vershynin (2012)**

$$\mathbb{E} \|\hat{\Sigma} - \Sigma\| \lesssim \max \left\{ \|\Sigma\|^{1/2} \mathbb{E}^{1/2} \max_{1 \leq j \leq n} \|X_j\|^2 \sqrt{\frac{\log d}{n}}, \mathbb{E} \max_{1 \leq j \leq n} \|X_j\|^2 \frac{\log d}{n} \right\}.$$

The proof is based on the approach by **Rudelson (1999)** and relies on noncommutative Khintchine inequality due to **Lust-Picard and Pisier (1991)**.

- Note that, in the subgaussian case,  $\left\| \|X\|^2 \right\|_{\psi_1} \lesssim \text{tr}(\Sigma)$ , which implies that

$$\mathbb{E} \max_{1 \leq j \leq n} \|X_j\|^2 \lesssim \text{tr}(\Sigma) \log n = \|\Sigma\| \mathbf{r}(\Sigma) \log n.$$

This implies

$$\mathbb{E} \|\hat{\Sigma} - \Sigma\| \lesssim \|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma) \log d \log n}{n}}, \frac{\mathbf{r}(\Sigma) \log d \log n}{n} \right\}.$$

If  $X$  is subgaussian, then with some constant  $C > 0$  and with probability at least  $1 - e^{-t}$

$$\|\hat{\Sigma} - \Sigma\| \leq C\|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma) \log d + t}{n}}, \frac{\mathbf{r}(\Sigma) \log d + t \log n}{n} \right\}.$$

The proof is based on a version of noncommutative Bernstein type inequality ([Ahlsvede and Winter \(2002\)](#), [Tropp \(2012\)](#)).

# Bounds for subgaussian r.v. in a separable Banach space

- $E$  a separable Banach space
- Recall that

$$\mathbf{r}(\Sigma) = \frac{\mathbb{E}\|Y\|^2}{\|\Sigma\|}, \quad Y \sim N(0; \Sigma)$$

## Definition

A weakly square integrable centered random variable  $X$  in  $E$  with covariance operator  $\Sigma$  is called **pregaussian** iff there exists a centered Gaussian random variable  $Y$  in  $E$  with the same covariance operator  $\Sigma$ .

# Bounds for subgaussian r.v. in a separable Banach space (Koltchinskii and Lounici (2014))

## Theorem

Let  $X, X_1, \dots, X_n$  be i.i.d. weakly square integrable centered random vectors in  $E$  with covariance operator  $\Sigma$ . If  $X$  is subgaussian and pregaussian, then

$$\mathbb{E}\|\hat{\Sigma} - \Sigma\| \lesssim \|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma)}{n}}, \frac{\mathbf{r}(\Sigma)}{n} \right\}$$

Moreover, if  $X$  is Gaussian, then

$$\|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma)}{n}}, \frac{\mathbf{r}(\Sigma)}{n} \right\} \lesssim \mathbb{E}\|\hat{\Sigma} - \Sigma\| \lesssim \|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma)}{n}}, \frac{\mathbf{r}(\Sigma)}{n} \right\}.$$



## Lemma (Decoupling)

Let  $X_1, \dots, X_n, X'_1, \dots, X'_n$  be i.i.d.  $N(0; \Sigma)$ . Then

$$\mathbb{E} \|\hat{\Sigma} - \Sigma\| \leq 2\mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle \langle X'_j, v \rangle \right|.$$

# Proof of the Upper Bound: Gaussian Case

## Proof.

$$\begin{aligned}\mathbb{E}\|\hat{\Sigma} - \Sigma\| &= \mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle \langle X_j, v \rangle - \mathbb{E} \langle X, u \rangle \langle X, v \rangle \right| \\ &= \mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \left| \mathbb{E}' n^{-1} \sum_{j=1}^n \left( \langle X_j, u \rangle \langle X_j, v \rangle + \langle X'_j, u \rangle \langle X_j, v \rangle \right. \right. \\ &\quad \left. \left. - \langle X_j, u \rangle \langle X'_j, v \rangle - \langle X'_j, u \rangle \langle X'_j, v \rangle \right) \right| \\ &\leq 2\mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \left\langle \frac{X_j + X'_j}{\sqrt{2}}, u \right\rangle \left\langle \frac{X_j - X'_j}{\sqrt{2}}, v \right\rangle \right| \\ &= 2\mathbb{E} \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle \langle X'_j, v \rangle \right|.\end{aligned}$$

# Proof of the Upper Bound: Gaussian Case



$$Y(u, v) := n^{-1/2} \sum_{j=1}^n \langle X_j, u \rangle \langle X'_j, v \rangle$$



$$Z(u, v) := \sqrt{2} \|\hat{\Sigma}'\|^{1/2} \langle X, u \rangle + \sqrt{2} \|\Sigma\|^{1/2} \left\langle \frac{1}{\sqrt{n}} \sum_{j=1}^n g_j X'_j, v \right\rangle,$$

where  $\hat{\Sigma}$  is the sample covariance based on  $X'_1, \dots, X'_n$  and  $\{g_j\}$  are i.i.d,  $N(0, 1)$  r.v. independent of  $\{X_j\}, \{X'_j\}$

- Conditionally on  $X'_j, j = 1, \dots, n$ ,  $(u, v) \mapsto Y(u, v)$  and  $(u, v) \mapsto Z(u, v)$  are mean zero Gaussian processes

Gaussian Comparison Inequality (Slepian-Fernique-Sudakov):  
conditionally on  $X'_1, \dots, X'_n$ ,

- $$\mathbb{E}_{X,g}(Y(u, v) - Y(u', v'))^2 \leq \mathbb{E}_{X,g}(Z(u, v) - Z(u', v'))^2.$$

- $$\begin{aligned} \mathbb{E}_{X,g} \sup_{\|u\|, \|v\| \leq 1} Y(u, v) &\leq \mathbb{E}_{X,g} \sup_{\|u\|, \|v\| \leq 1} Z(u, v) \\ &\leq \sqrt{2} \|\hat{\Sigma}'\|^{1/2} \mathbb{E} \|X\| + \sqrt{2} \|\Sigma\|^{1/2} \mathbb{E}_g \left\| \frac{1}{\sqrt{n}} \sum_{j=1}^n g_j X'_j \right\|. \end{aligned}$$

# Proof of the Upper Bound: Gaussian Case

- Combining this with the decoupling inequality and using the fact that

$$\mathbb{E} \left\| \frac{1}{\sqrt{n}} \sum_{j=1}^n g_j X_j' \right\| = \mathbb{E} \left( n^{-1} \sum_{j=1}^n g_j^2 \right)^{1/2} \mathbb{E} \|X\| \leq \mathbb{E} \|X\|,$$

- we get

$$\begin{aligned} \Delta &:= \mathbb{E} \|\hat{\Sigma} - \Sigma\| \\ &\leq 2\sqrt{2}\Delta^{1/2} \frac{\mathbb{E} \|X\|}{\sqrt{n}} + 4\sqrt{2} \|\Sigma\|^{1/2} \frac{\mathbb{E} \|X\|}{\sqrt{n}} \\ &\leq 2\sqrt{2}\Delta^{1/2} \|\Sigma\|^{1/2} \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} + 4\sqrt{2} \|\Sigma\| \sqrt{\frac{\mathbf{r}(\Sigma)}{n}}, \end{aligned}$$

and the upper bound follows by solving the above inequality w.r.t.  $\Delta$ .

# Proof of the Upper Bound: Subgaussian Case

The proof is based on generic chaining bounds for empirical processes indexed by squares of functions ([Mendelson \(2012\)](#))

- $(T, d)$  a metric space
- $\{\Delta_n\}$  an increasing sequence of partitions of  $T$
- $\{\Delta_n\}$  admissible iff  $\text{card}(\Delta_n) \leq N_n$ , where  $N_n := 2^{2^n}$ ,  $n \geq 1$ ,  $N_0 := 1$ .
- For  $t \in T$ ,  $\Delta_n(t)$  denotes the unique set of the partition  $\Delta_n$  that contains  $t$ .
- $A \subset T$ ,  $D(A)$  denotes the diameter of set  $A$ .
- **Generic Chaining Complexity**

$$\gamma_2(T, d) = \inf \sup_{t \in T} \sum_{n=0}^{\infty} 2^{n/2} D(\Delta_n(t)),$$

where the infimum is taken over all admissible sequences.

## Theorem

Let  $X(t), t \in T$  be a centered Gaussian process and suppose that

$$d(t, s) := \mathbb{E}^{1/2}(X(t) - X(s))^2, t, s \in T.$$

Then, there exists an absolute constant  $K > 0$  such that

$$K^{-1} \gamma_2(T; d) \leq \mathbb{E} \sup_{t \in T} X(t) \leq K \gamma_2(T; d)$$



## Theorem

Let  $X, X_1, \dots, X_n$  be i.i.d. random variables in  $S$  with common distribution  $P$  and let  $\mathcal{F}$  be a class of measurable functions on  $(S, \mathcal{A})$  such that  $f \in \mathcal{F}$  implies  $-f \in \mathcal{F}$  and  $\mathbb{E}f(X) = 0$ . Then

$$\mathbb{E} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n f^2(X_i) - \mathbb{E}f^2(X) \right| \lesssim \max \left\{ \sup_{f \in \mathcal{F}} \|f\|_{\psi_1} \frac{\gamma_2(\mathcal{F}; \psi_2)}{\sqrt{n}}, \frac{\gamma_2^2(\mathcal{F}; \psi_2)}{n} \right\}.$$

# Proof of the Upper Bound: Subgaussian Case



$$\begin{aligned}\mathbb{E} \|\hat{\Sigma} - \Sigma\| &= \mathbb{E} \sup_{\|u\| \leq 1} \left| \frac{1}{n} \sum_{i=1}^n \langle X_i, u \rangle^2 - \langle \Sigma u, u \rangle \right| \\ &= \mathbb{E} \sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n f^2(X_i) - \mathbb{E} f^2(X) \right|,\end{aligned}$$

where  $\mathcal{F} := \{ \langle \cdot, u \rangle : u \in U_{E^*} \}$ ,  $U_{E^*} := \{ u \in E^* : \|u\| \leq 1 \}$  and  $P$  is the distribution of random variable  $X$ .

- Since  $X$  is subgaussian,

$$\left\| \langle X, u \rangle \right\|_{\psi_2} \asymp \left\| \langle X, u \rangle \right\|_{\psi_1} \asymp \left\| \langle X, u \rangle \right\|_{L_2(\mathbb{P})}.$$

- Therefore,

$$\sup_{f \in \mathcal{F}} \|f\|_{\psi_1} \lesssim \sup_{u \in U_{E^*}} \mathbb{E}^{1/2} \langle X, u \rangle^2 \leq \|\Sigma\|^{1/2}.$$

# Proof of the Upper Bound: Subgaussian Case

- Also, since  $X$  is pregaussian, there exists  $Y \sim N(0, \Sigma)$
- 

$$d_Y(u, v) = \|\langle \cdot, u \rangle - \langle \cdot, v \rangle\|_{L_2(P)}, u, v \in U_{E^*}.$$

- Using Talagrand Theorem,

$$\gamma_2(\mathcal{F}, \psi_2) \lesssim \gamma_2(\mathcal{F}, L_2) = \gamma_2(U_{E^*}; d_Y) \lesssim \mathbb{E} \sup_{u \in U_{E^*}} \langle Y, u \rangle \leq \mathbb{E} \|Y\|.$$

Therefore,

$$\begin{aligned} \mathbb{E} \|\hat{\Sigma} - \Sigma\| &\lesssim \max \left\{ \|\Sigma\|^{1/2} \frac{\mathbb{E} \|Y\|}{\sqrt{n}}, \frac{(\mathbb{E} \|Y\|)^2}{n} \right\} \\ &\lesssim \|\Sigma\| \max \left\{ \sqrt{\frac{\mathbf{r}(\Sigma)}{n}}, \frac{\mathbf{r}(\Sigma)}{n} \right\}. \end{aligned}$$

# Proof of the Lower Bound



$$\mathbb{E} \|\hat{\Sigma} - \Sigma\| \geq \sup_{\|u\| \leq 1} \mathbb{E} \left\| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j - \mathbb{E} \langle X, u \rangle X \right\|.$$

- For a fixed  $u \in E^*$  with  $\|u\| \leq 1$  and  $\langle \Sigma u, u \rangle > 0$ , define

$$X' := X - \langle X, u \rangle \frac{\Sigma u}{\langle \Sigma u, u \rangle}, \quad X'_j := X_j - \langle X_j, u \rangle \frac{\Sigma u}{\langle \Sigma u, u \rangle}, j = 1, \dots, n.$$

$\{X', X'_j : j = 1, \dots, n\}$  and  $\{\langle X, u \rangle, \langle X_j, u \rangle : j = 1, \dots, n\}$  are independent.



$$\begin{aligned} \mathbb{E} \left\| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j - \mathbb{E} \langle X, u \rangle X \right\| &= \\ \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\Sigma u}{\langle \Sigma u, u \rangle} + n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X'_j \right\|, \end{aligned}$$

# Proof of the Lower Bound

- Conditionally on  $\langle X_j, u \rangle, j = 1, \dots, n$ , the distribution of r.v.

$$n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j'$$

is Gaussian and it coincides with the distribution of r.v.

$$\left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}}.$$



$$\begin{aligned} & \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\sum u}{\langle \sum u, u \rangle} + n^{-1} \sum_{j=1}^n \langle X_j, u \rangle X_j' \right\| \\ &= \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\sum u}{\langle \sum u, u \rangle} + \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\|. \end{aligned}$$

# Proof of the Lower Bound

Denote  $\mathbb{E}'$  the conditional expectation given  $X'_1, \dots, X'_n$ .

$$\begin{aligned} & \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\sum u}{\langle \sum u, u \rangle} + \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\| \\ &= \mathbb{E} \mathbb{E}' \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\sum u}{\langle \sum u, u \rangle} + \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\| \\ &\geq \mathbb{E} \left\| \mathbb{E}' n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\sum u}{\langle \sum u, u \rangle} + \mathbb{E}' \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\| \\ &= \mathbb{E} \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{\mathbb{E} \|X'\|}{\sqrt{n}}. \end{aligned}$$

# Proof of the Lower Bound

$$\mathbb{E}\|X'\| \geq \mathbb{E}\|X\| - \mathbb{E}|\langle X, u \rangle| \frac{\|\Sigma u\|}{\langle \Sigma u, u \rangle} = \mathbb{E}\|X\| - \sqrt{\frac{2}{\pi}} \frac{\|\Sigma u\|}{\langle \Sigma u, u \rangle^{1/2}}$$

and

$$\begin{aligned} \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E}\langle X, u \rangle^2) \frac{\Sigma u}{\langle \Sigma u, u \rangle} + \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\| \\ \geq \langle \Sigma u, u \rangle^{1/2} \mathbb{E} \left( n^{-1} \sum_{j=1}^n Z_j^2 \right)^{1/2} \frac{\mathbb{E}\|X\| - \sqrt{\frac{2}{\pi}} \frac{\|\Sigma u\|}{\langle \Sigma u, u \rangle^{1/2}}}{\sqrt{n}}, \end{aligned}$$

where

$$Z_j = \frac{\langle X_j, u \rangle}{\langle \Sigma u, u \rangle^{1/2}}, j = 1, \dots, n \text{ i.i.d. } \sim N(0, 1)$$

# Proof of the Lower Bound

Since

$$\mathbb{E} \left( n^{-1} \sum_{j=1}^n Z_j^2 \right)^{1/2} \geq c_2 > 0,$$

$$\begin{aligned} \mathbb{E} \left\| n^{-1} \sum_{j=1}^n (\langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2) \frac{\Sigma u}{\langle \Sigma u, u \rangle} + \left( n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 \right)^{1/2} \frac{X'}{\sqrt{n}} \right\| \\ \geq c_2 \frac{\langle \Sigma u, u \rangle^{1/2} \mathbb{E} \|X\| - \sqrt{\frac{2}{\pi}} \|\Sigma u\|}{\sqrt{n}}. \end{aligned}$$

Therefore

$$\begin{aligned} \mathbb{E} \|\hat{\Sigma} - \Sigma\| &\geq c_2 \sup_{\|u\| \leq 1} \frac{\langle \Sigma u, u \rangle^{1/2} \mathbb{E} \|X\| - \sqrt{\frac{2}{\pi}} \|\Sigma u\|}{\sqrt{n}} \\ &\geq c_2 \frac{\|\Sigma\|^{1/2} \mathbb{E} \|X\| - \sqrt{\frac{2}{\pi}} \|\Sigma\|}{\sqrt{n}} \geq c_2 \|\Sigma\| \left( \frac{c_3 \sqrt{\mathbf{r}(\Sigma)} - \sqrt{\frac{2}{\pi}}}{\sqrt{n}} \right). \end{aligned}$$



# Proof of the Lower Bound

Also,

$$\begin{aligned}\mathbb{E}\|\hat{\Sigma} - \Sigma\| &\geq \sup_{\|u\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 - \mathbb{E} \langle X, u \rangle^2 \right| \\ &= \sup_{\|u\| \leq 1} \langle \Sigma u, u \rangle \mathbb{E} \left| n^{-1} \sum_{j=1}^n Z_j^2 - 1 \right| \geq c_4 \frac{\|\Sigma\|}{\sqrt{n}},\end{aligned}$$

implying that, for small enough  $c_2$ ,

$$\begin{aligned}\mathbb{E}\|\hat{\Sigma} - \Sigma\| &\geq c_2 \|\Sigma\| \left( \frac{c_3 \sqrt{\mathbf{r}(\Sigma)} - \sqrt{\frac{2}{\pi}}}{\sqrt{n}} \right) \vee c_4 \frac{\|\Sigma\|}{\sqrt{n}} \\ &\geq \frac{1}{2} \left( c_2 \|\Sigma\| \left( \frac{c_3 \sqrt{\mathbf{r}(\Sigma)} - \sqrt{\frac{2}{\pi}}}{\sqrt{n}} \right) + c_4 \frac{\|\Sigma\|}{\sqrt{n}} \right) \geq \frac{c_2}{2} \|\Sigma\| \frac{c_3 \sqrt{\mathbf{r}(\Sigma)}}{\sqrt{n}}.\end{aligned}$$

# Proof of the Lower Bound

On the other hand, if  $\mathbf{r}(\Sigma) \geq 2n$ ,

$$\begin{aligned}\mathbb{E}\|\hat{\Sigma} - \Sigma\| &\geq \mathbb{E}\|\hat{\Sigma}\| - \|\Sigma\| \geq \mathbb{E} \sup_{\|u\| \leq 1} n^{-1} \sum_{j=1}^n \langle X_j, u \rangle^2 - \|\Sigma\| \\ &\geq \mathbb{E} \sup_{\|u\| \leq 1} \frac{\langle X_1, u \rangle^2}{n} - \|\Sigma\| \geq \frac{\mathbb{E}\|X\|^2}{n} - \|\Sigma\| \\ &= \|\Sigma\| \left( \frac{\mathbf{r}(\Sigma)}{n} - 1 \right) \geq \frac{1}{2} \|\Sigma\| \frac{\mathbf{r}(\Sigma)}{n}.\end{aligned}$$

# Concentration inequality (Koltchinskii and Lounici (2014))

## Theorem

Let  $M$  be either the median, or the mean of  $\|\hat{\Sigma} - \Sigma\|$ . There exists a constant  $C > 0$  such that, for all  $t \geq 1$  with probability at least  $1 - e^{-t}$ , the following bound holds:

$$\left| \|\hat{\Sigma} - \Sigma\| - M \right| \leq C \left[ \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \sqrt{\frac{t}{n}} \vee \|\Sigma\| \frac{t}{n} \right].$$

## Corollary

*There exists a constant  $C > 0$  such that, for all  $t \geq 1$ , with probability at least  $1 - e^{-t}$ ,*

$$\|\hat{\Sigma} - \Sigma\| \leq C\|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \vee \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right).$$

*This implies that for all  $p \geq 1$*

$$\mathbb{E}^{1/p} \|\hat{\Sigma} - \Sigma\|^p \lesssim_p \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right).$$

# Proof of concentration inequality:

- The proof is based on Gaussian concentration
- Another proof: based on a concentration inequality for sup-norms of Gaussian chaos, [Adamczak \(2014\)](#)

## Theorem

Let  $X, X_1, \dots, X_n$  be i.i.d. centered Gaussian random vectors in  $E$  with covariance  $\Sigma$  and let  $M := \text{Med}(\|\hat{\Sigma} - \Sigma\|)$ . Then, there exist constants  $C > 0$  such that for all  $t \geq 1$  with probability at least  $1 - e^{-t}$ ,

$$\left| \|\hat{\Sigma} - \Sigma\| - M \right| \leq C \left[ \|\Sigma\| \left( \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right) \vee \|\Sigma\|^{1/2} M^{1/2} \sqrt{\frac{t}{n}} \right].$$

Note that

$$M = \text{Med}(\|\hat{\Sigma} - \Sigma\|) \leq 2\mathbb{E}\|\hat{\Sigma} - \Sigma\| \lesssim \|\Sigma\| \left[ \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right],$$

implying the concentration inequality in the explicit form.

# Proof of Concentration Inequality, reduction to the finite-dimensional case

## Theorem

Let  $X$  be a centered Gaussian random variable in a separable Banach space  $E$ . Then there exists a sequence  $\{x_k : k \geq 1\}$  of vectors in  $E$  and a sequence  $\{Z_k : k \geq 1\}$  of i.i.d. standard normal random variables such that  $X = \sum_{k=1}^{\infty} Z_k x_k$ , where the series in the right hand side converges in  $E$  a.s. and  $\sum_{k=1}^{\infty} \|x_k\|^2 < +\infty$ .

It easily follows from this result that it is enough to proof the concentration inequality when

$$X = \sum_{k=1}^m Z_k x_k, X_j = \sum_{k=1}^m Z_{k,j} x_k, Z := (Z_{k,j}, k = 1, \dots, m, j = 1, \dots, n).$$

# Gaussian Concentration

Denote

$$f(Z) =: g(X_1, \dots, X_n) := \|W\|_\varphi\left(\frac{\|W\|}{\delta}\right),$$

where

- $W = \hat{\Sigma} - \Sigma$ ,
- $\varphi$  is a Lipschitz function with constant 1 on  $\mathbb{R}_+$ ,  $0 \leq \varphi(s) \leq 1$ ,  
 $\varphi(s) = 1, s \leq 1, \varphi(s) = 0, s > 2$ ,
- $\delta > 0$  is fixed

## Lemma

There exists a numerical constant  $D > 0$  such that, for all  $Z, Z' \in \mathbb{R}^{mn}$ ,

$$|f(Z) - f(Z')| \leq D \frac{\|\Sigma\| + \|\Sigma\|^{1/2} \sqrt{\delta}}{\sqrt{n}} \left( \sum_{j=1}^n \sum_{k=1}^m |Z_{k,j} - Z'_{k,j}|^2 \right)^{1/2}.$$



# Gaussian Concentration

By Gaussian concentration inequality, for all  $t \geq 1$  with probability at least  $1 - e^{-t}$ ,

$$\left| g(X_1, \dots, X_n) - \text{Med}(g(X_1, \dots, X_n)) \right| \leq D_1 \left( \|\Sigma\| + \|\Sigma\|^{1/2} \sqrt{\delta} \right) \sqrt{\frac{t}{n}},$$

where  $D_1$  is a numerical constant. It follows that, on the event  $\|W\| \leq \delta$ ,

$$\begin{aligned} \|W\| = g(X_1, \dots, X_n) &\leq \text{Med}(g(X_1, \dots, X_n)) + D_1 \left( \|\Sigma\| + \|\Sigma\|^{1/2} \sqrt{\delta} \right) \sqrt{\frac{t}{n}} \\ &\leq \text{Med}(\|W\|) + D_1 \left( \|\Sigma\| + \|\Sigma\|^{1/2} \sqrt{\delta} \right) \sqrt{\frac{t}{n}} =: A + B\sqrt{\delta}, \end{aligned}$$

where

$$A := \text{Med}(\|W\|) + D_1 \|\Sigma\| \sqrt{\frac{t}{n}}, \quad B := D_1 \|\Sigma\|^{1/2} \sqrt{\frac{t}{n}}.$$

Then we have

$$\mathbb{P} \left\{ \delta \geq \|W\| \geq A + B\sqrt{\delta} \right\} \leq e^{-t}.$$

# Proof of Concentration Inequality: Iterative Bounds

- Denote

$$\delta_0 := D_2 \|\Sigma\| \left[ \mathbf{r}(\Sigma) \left( \sqrt{\frac{t}{n}} V \frac{t}{n} \right) + \mathbf{r}(\Sigma) + 1 \right].$$

- It is easy to prove that

$$\mathbb{P} \left\{ \|W\| \geq \delta_0 \right\} \leq e^{-t}$$

- Define  $\delta_k$  for  $k \geq 1$  as follows:

$$\delta_k = A + B\sqrt{\delta_{k-1}}.$$

- It is easy to check that  $\{\delta_k\}$  is decreasing with limit  $\bar{\delta}$ ,

$$\bar{\delta} = A + B\sqrt{\bar{\delta}}, \quad \bar{\delta} \lesssim A \vee B^2.$$

- Moreover,

$$\delta_k - \bar{\delta} \leq u_k := B^2 \left( \frac{\delta_0}{B^2} \right)^{2^{-k}}.$$

# Proof of Concentration Inequality: Iterative Bounds

- Let  $\bar{k} := \min \left\{ k : \left( \frac{\delta_0}{B^2} \right)^{2^{-k}} \leq 2 \right\}$ . Then

$$\delta_{\bar{k}} \lesssim A \vee B^2, \quad \bar{k} \lesssim \log \log(c_1 \mathbf{r}(\Sigma)) \vee \log \log(c_1 n)$$

- Since  $\mathbb{P} \left\{ \delta_{k-1} > \|W\| \geq \delta_k \right\} \leq e^{-t}$ , we get that with probability at least  $1 - (\bar{k} + 1)e^{-t}$ ,

$$\|W\| \leq \delta_{\bar{k}} \lesssim A \vee B^2 \lesssim \text{Med}(\|W\|) \vee \|\Sigma\| \left( \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right)$$

- With a bit more effort, it follows that with probability at least  $1 - e^{-t}$

$$\|W\| = \|\hat{\Sigma} - \Sigma\| \leq C \left[ \text{Med}(\|W\|) \vee \|\Sigma\| \left( \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right) \right].$$

# Gaussian Concentration

- Take

$$\delta := C \left[ \text{Med}(\|W\|) \vee \|\Sigma\| \left( \sqrt{\frac{t+2n}{n}} \vee \frac{t+2n}{n} \right) \right].$$

- Since  $\mathbb{P}\{\|W\| \geq \delta\} \leq 2e^{-t-2n} \leq 1/4$ ,

$$\mathbb{P}\{g(X_1, \dots, X_n) \geq \text{Med}(\|W\|)\} \geq 1/4,$$

$$\mathbb{P}\{g(X_1, \dots, X_n) \leq \text{Med}(\|W\|)\} \geq 1/2.$$

- To complete the proof, note that, by Gaussian isoperimetric inequality, on the event where  $\|W\| \leq \delta$ ,

$$\begin{aligned} \left| \|\hat{\Sigma} - \Sigma\| - \text{Med}(\|W\|) \right| &= \left| g(X_1, \dots, X_n) - \text{Med}(\|W\|) \right| \\ &\leq D_1 \left( \|\Sigma\| + \|\Sigma\|^{1/2} \sqrt{\delta} \right) \sqrt{\frac{t}{n}} \end{aligned}$$

with probability at least  $1 - e^{-t}$ .

# Generic Chaining Tail Bound: Dirksen (2014), Bednorz (2014), Mendelson (2013, 2015)

## Theorem

Let  $X, X_1, \dots, X_n$  be i.i.d. random variables in  $S$  with common distribution  $P$  and let  $\mathcal{F}$  be a class of measurable functions on  $(S, \mathcal{A})$ . Then, there exists a constant  $C > 0$  such that for all  $t \geq 1$  with probability at least  $1 - e^{-t}$

$$\sup_{f \in \mathcal{F}} \left| \frac{1}{n} \sum_{i=1}^n f^2(X_i) - \mathbb{E} f^2(X) \right| \leq C \max \left\{ \sup_{f \in \mathcal{F}} \|f\|_{\psi_2} \frac{\gamma_2(\mathcal{F}; \psi_2)}{\sqrt{n}}, \frac{\gamma_2^2(\mathcal{F}; \psi_2)}{n}, \sup_{f \in \mathcal{F}} \|f\|_{\psi_2}^2 \sqrt{\frac{t}{n}}, \sup_{f \in \mathcal{F}} \|f\|_{\psi_2}^2 \frac{t}{n} \right\}.$$

## Part 2.

# Taylor Expansions of Operator Functions and Normal Approximation of Plug-In Estimators of Smooth Functionals of Covariance

- Let  $\mathbb{H}$  be a separable Hilbert space
- Let  $X, X_1, \dots, X_n$  be i.i.d. Gaussian vectors with values in  $\mathbb{H}$  with  $\mathbb{E}X = 0$  and with covariance operator  $\Sigma = \mathbb{E}(X \otimes X)$ .
- **Problems**
  - Given a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$  and a nuclear operator  $B : \mathbb{H} \mapsto \mathbb{H}$ , estimate  $\langle f(\Sigma), B \rangle$  based on  $X_1, \dots, X_n$ .

# Sample Covariance Operator and Effective Rank

- Let

$$\hat{\Sigma} := n^{-1} \sum_{j=1}^n X_j \otimes X_j$$

be the sample covariance based on  $(X_1, \dots, X_n)$ .

- Effective Rank:**

$$\mathbf{r}(\Sigma) = \frac{\text{tr}(\Sigma)}{\|\Sigma\|}$$

- $\mathbf{r}(\Sigma) \leq \text{rank}(\Sigma) \leq \dim(\mathbb{H})$



## Theorem (Koltchinskii and Lounici (2014))

Let  $X, X_1, \dots, X_n$  be i.i.d centered Gaussian random vectors in  $\mathbb{H}$  with covariance operator  $\Sigma$ . Then

$$\mathbb{E}\|\hat{\Sigma} - \Sigma\| \asymp \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \right).$$

# Concentration Inequality

## Theorem (Koltchinskii and Lounici (2014))

There exists a constant  $C > 0$  such that, for all  $t \geq 1$  with probability at least  $1 - e^{-t}$ , the following bound holds:

$$\left| \|\hat{\Sigma} - \Sigma\| - M \right| \leq C \|\Sigma\| \left[ \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right].$$

where  $M$  is either the mean, or the median of  $\|\hat{\Sigma} - \Sigma\|$ .

## Remark

The results are also true for Gaussian random variables in separable Banach spaces with  $\mathbf{r}(\Sigma) := \frac{\mathbb{E}\|X\|^2}{\|\Sigma\|}$ .

## Corollary

There exists a constant  $C > 0$  such that, for all  $t \geq 1$ , with probability at least  $1 - e^{-t}$ ,

$$\|\hat{\Sigma} - \Sigma\| \leq C \|\Sigma\| \left( \sqrt{\frac{r(\Sigma)}{n}} \vee \frac{r(\Sigma)}{n} \vee \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right).$$

This implies that for all  $p \geq 1$

$$\mathbb{E}^{1/p} \|\hat{\Sigma} - \Sigma\|^p \lesssim_p \|\Sigma\| \left( \sqrt{\frac{r(\Sigma)}{n}} \vee \frac{r(\Sigma)}{n} \right).$$

# Normal Approximation Bounds for Smooth Functions of Sample Covariance

## Problems

Let  $f : \mathbb{R} \mapsto \mathbb{R}$  be a given smooth function and  $B$  be a given operator with  $\|B\|_1 < \infty$

- For given  $f$  and  $B$ , show that the distribution of r.v.  $\frac{n^{1/2} \langle f(\hat{\Sigma}) - \mathbb{E}f(\hat{\Sigma}), B \rangle}{\sigma_f(\Sigma; B)}$  is close to standard normal for a proper  $\sigma_f(\Sigma; B)$  when  $n \rightarrow \infty$  and  $\mathbf{r}(\Sigma) = o(n)$
- Is the plug-in estimator  $\langle f(\hat{\Sigma}), B \rangle$  asymptotically efficient (is  $\sqrt{n} \langle f(\hat{\Sigma}) - f(\Sigma), B \rangle$  asymptotically normal with limit variance  $\sigma_f(\Sigma; B)$  as small as possible)?

# Entire Functions of Exponential Type

- $f : \mathbb{C} \mapsto \mathbb{C}$  be an entire function
- For  $\sigma > 0$ ,  $f$  is of **exponential type  $\sigma$**  if  $\forall \varepsilon > 0 \exists C = C(\varepsilon, \sigma, f) > 0$  such that

$$|f(z)| \leq C e^{(\sigma + \varepsilon)|z|}, z \in \mathbb{C}.$$

- $\mathcal{E}_\sigma = \mathcal{E}_\sigma(\mathbb{C})$  denotes the space of all entire functions of exponential type  $\sigma$ .
- According to *Paley-Wiener theorem*,

$$\mathcal{E}_\sigma \cap L_\infty(\mathbb{R}) = \{f \in L_\infty(\mathbb{R}) : \text{supp}(\mathcal{F}f) \subset [-\sigma, \sigma]\}.$$

- *Bernstein inequality*:  $\forall f \in \mathcal{E}_\sigma \cap L_\infty(\mathbb{R})$

$$\|f'\|_{L_\infty(\mathbb{R})} \leq \sigma \|f\|_{L_\infty(\mathbb{R})}.$$

# Littlewood-Paley Decomposition

- Let  $w \in C^\infty(\mathbb{R})$ ,  $w \geq 0$ ,  $\text{supp}(w) \subset [-2, 2]$ ,  $w(t) = 1$ ,  $t \in [-1, 1]$  and  $w(-t) = w(t)$ ,  $t \in \mathbb{R}$ .
- $w_0(t) := w(t/2) - w(t)$ ,  $t \in \mathbb{R}$ ,  $\text{supp}(w_0) \subset \{t : 1 \leq |t| \leq 4\}$
- $w_j(t) := w_0(2^{-j}t)$ ,  $t \in \mathbb{R}$ ,  $\text{supp}(w_j) \subset \{t : 2^j \leq |t| \leq 2^{j+2}\}$ ,  $j \geq 0$ .
- Then  $w(t) + \sum_{j \geq 0} w_j(t) = 1$ ,  $t \in \mathbb{R}$ .
- Let  $W, W_j \in \mathcal{S}(\mathbb{R})$ ,

$$w(t) = (\mathcal{F}W)(t), \quad w_j(t) = (\mathcal{F}W_j)(t), \quad t \in \mathbb{R}, j \geq 0.$$

- For  $f \in \mathcal{S}'(\mathbb{R})$ , define its **Littlewood-Paley dyadic decomposition**:

$$f_0 := f * W, \quad f_n := f * W_{n-1}, \quad n \geq 1$$

- Note that  $f_n \in \mathcal{E}_{2^{n+1}} \cap L_\infty(\mathbb{R})$  and

$$\sum_{n \geq 0} f_n = f$$

with convergence of the series in the space  $\mathcal{S}'(\mathbb{R})$ .

- *Besov norms:*

$$\|f\|_{B_{\infty,1}^s} := \sum_{n \geq 0} 2^{ns} \|f_n\|_{L_{\infty}(\mathbb{R})}, \quad s \in \mathbb{R}$$

- *Besov spaces:*

$$B_{\infty,1}^s(\mathbb{R}) := \left\{ f \in \mathcal{S}'(\mathbb{R}) : \|f\|_{B_{\infty,1}^s} < \infty \right\}.$$

- If  $f \in B_{\infty,1}^s(\mathbb{R})$  for some  $s \geq 0$ , then  $\sum_{n \geq 0} f_n$  converges uniformly to  $f$  in  $\mathbb{R}$ , which easily implies that  $f \in \mathcal{C}_u(\mathbb{R})$  and

$$\|f\|_{L_{\infty}} \leq \|f\|_{B_{\infty,1}^s}.$$

# Perturbation Theory: Operator Lipschitz and Operator Differentiable Functions

- $\mathcal{B}_{sa}(\mathbb{H})$  the space of self-adjoint bounded operators in  $\mathbb{H}$
- A continuous function  $f : \mathbb{R} \mapsto \mathbb{R}$  is called **operator Lipschitz** with respect to the operator norm iff there exists a constant  $L_f > 0$  such that for all  $A, B \in \mathcal{B}_{sa}(\mathbb{H})$

$$\|f(A) - f(B)\| \leq L_f \|A - B\|.$$

- If  $f$  is operator Lipschitz, then it is Lipschitz; however,  $f(t) = |t|$  is not operator Lipschitz (**Kato (1972)**).
- A continuous function  $f : \mathbb{R} \mapsto \mathbb{R}$  is called **operator differentiable** iff  $\mathcal{B}_{sa}(\mathbb{H}) \ni A \mapsto f(A) \in \mathcal{B}_{sa}(\mathbb{H})$  is Fréchet differentiable at any  $A \in L(\mathbb{H})$ , i.e., there exists a bounded linear mapping  $\mathcal{B}_{sa}(\mathbb{H}) \ni E \mapsto Df(A; E) = Df(A)E \in \mathcal{B}_{sa}(\mathbb{H})$  such that

$$f(A + E) - f(A) = Df(A; E) + o(\|E\|) \text{ as } \|E\| \rightarrow 0.$$



# Perturbation Theory: Operator Lipschitz and Operator Differentiable Functions

## Theorem (Peller)

If  $f \in B_{\infty,1}^1(\mathbb{R})$  then  $f$  is operator Lipschitz with Lipschitz constant  $L_f = \|f\|_{B_{\infty,1}^1}$  and operator differentiable.

Moreover, if  $A \in \mathcal{B}_{sa}(\mathbb{H})$  is an operator with spectral decomposition

$$A = \sum_{\lambda \in \sigma(A)} \lambda P_{\lambda},$$

then (Loewner, Daletsky-Krein)

$$Df(A; E) = \sum_{\lambda, \mu \in \sigma(A)} f^{[1]}(\lambda, \mu) P_{\lambda} E P_{\mu},$$

where

$$f^{[1]}(\lambda, \mu) := \frac{f(\lambda) - f(\mu)}{\lambda - \mu}, \quad \lambda \neq \mu; \quad f^{[1]}(\lambda, \lambda) := f'(\lambda).$$

# Perturbation Theory: Bounds on the Remainder of Differentiation

## Lemma

Let  $S_f(A; E) = f(A + E) - f(A) - (Df)(A; E)$  be the remainder of differentiation. If, for some  $s \in [1, 2]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ , then the following bounds hold:

$$\|S_f(A; E)\| \leq 2^{3-s} \|f\|_{B_{\infty,1}^s} \|E\|^s$$

and

$$\|S_f(A; E) - S_f(A; E')\| \leq 2^{1+s} \|f\|_{B_{\infty,1}^s} (\|E\| \vee \|E'\|)^{s-1} \|E' - E\|.$$

The proof is based on Littlewood-Paley decomposition of  $f$  and on operator versions of Bernstein inequalities for entire functions of exponential type (as in the work by **Peller** on operator Lipschitz functions).

## Lemma

Let  $f \in \mathcal{E}_\sigma \cap L_\infty(\mathbb{R})$ . Then, for all  $A, H, H' \in \mathcal{B}_{sa}(\mathbb{H})$ ,

$$\|f(A+H) - f(A)\| \leq \sigma \|f\|_{L_\infty(\mathbb{R})} \|H\|,$$

$$\|Df(A; H)\| \leq \sigma \|f\|_{L_\infty(\mathbb{R})} \|H\|,$$

$$\|S_f(A; H)\| \leq \frac{\sigma^2}{2} \|f\|_{L_\infty(\mathbb{R})} \|H\|^2$$

and

$$\|S_f(A; H') - S_f(A; H)\| \leq \sigma^2 \|f\|_{L_\infty(\mathbb{R})} \delta(H, H') \|H' - H\|,$$

where

$$\delta(H, H') := (\|H\| \wedge \|H'\|) + \frac{\|H' - H\|}{2}.$$

# Proof of Operator Lipschitz Property

- $E$  a complex Banach space
- $\mathcal{E}_\sigma(E)$  the space of entire functions  $F : \mathbb{C} \mapsto E$  of exponential type  $\sigma$  :  $F \in \mathcal{E}_\sigma(E)$  iff  $\forall \varepsilon > 0 \exists C = C(\varepsilon, \sigma, F) > 0$  :

$$\|F(z)\| \leq C e^{(\sigma+\varepsilon)|z|}, z \in \mathbb{C}.$$

- If  $F \in \mathcal{E}_\sigma(E)$  and  $\sup_{x \in \mathbb{R}} \|F(x)\| < +\infty$ , then Bernstein inequality holds for  $F$  :

$$\sup_{x \in \mathbb{R}} \|F'(x)\| \leq \sigma \sup_{x \in \mathbb{R}} \|F(x)\|$$

and

$$\|F(x+h) - F(x)\| \leq \sigma \sup_{x \in \mathbb{R}} \|F(x)\| |h|.$$

# Proof of Operator Lipschitz Property

- Given  $A, H \in \mathcal{B}_{sa}(\mathbb{H})$  and  $f \in \mathcal{E}_\sigma \cap L_\infty(\mathbb{R})$ , define

$$F(z) := f(A + zH), z \in \mathbb{C}.$$

- Then  $F \in \mathcal{E}_{\sigma\|H\|}(\mathcal{B}(\mathbb{H}))$ . Indeed,  $F$  is complex-differentiable at any point  $z \in \mathbb{C}$  with derivative  $F'(z) = Df(A + zH; H)$  and, by von Neumann theorem,

$$\|F(z)\| = \|f(A + zH)\| \leq \sup_{|\zeta| \leq \|A\| + |z|\|H\|} |f(\zeta)| \leq \|f\|_{L_\infty(\mathbb{R})} e^{\sigma\|A\|} e^{\sigma\|H\||z|},$$

implying that  $F$  is of exponential type  $\sigma\|H\|$ .

# Proof of Operator Lipschitz Property

- Note also that

$$\sup_{x \in \mathbb{R}} \|F(x)\| = \sup_{x \in \mathbb{R}} \|f(A + xH)\| \leq \sup_{x \in \mathbb{R}} |f(x)| = \|f\|_{L_\infty(\mathbb{R})}.$$

- Hence

$$\begin{aligned} \|f(A + H) - f(A)\| &= \|F(1) - F(0)\| \leq \sup_{x \in \mathbb{R}} \|F'(x)\| \\ &\leq \sigma \|H\| \sup_{x \in \mathbb{R}} \|F(x)\| \leq \sigma \|f\|_{L_\infty(\mathbb{R})} \|H\|. \end{aligned}$$

# Proof Operator Lipschitz Property for $f \in B_{\infty,1}^1(\mathbb{R})$

- For  $f \in B_{\infty,1}^1(\mathbb{R})$ , the series  $\sum_{n \geq 0} f_n$  converges uniformly in  $\mathbb{R}$  to function  $f$ .
- Since  $A, A + H, A + H'$  are bounded self-adjoint operators, we also get

$$\sum_{n \geq 0} f_n(A) = f(A), \quad \sum_{n \geq 0} f_n(A+H) = f(A+H), \quad \sum_{n \geq 0} f_n(A+H') = f(A+H')$$

with convergence of the series in the operator norm.



$$\begin{aligned} \|f(A+H) - f(A)\| &= \left\| \sum_{n \geq 0} [f_n(A+H) - f_n(A)] \right\| \\ &\leq \sum_{n \geq 0} \|f_n(A+H) - f_n(A)\| \leq \sum_{n \geq 0} 2^{n+1} \|f_n\|_{L^\infty(\mathbb{R})} \|H\| = 2 \|f\|_{B_{\infty,1}^1} \|H\|. \end{aligned}$$

# Higher Order Operator Differentiability

- If  $g : \mathcal{B}_{sa}(\mathbb{H}) \mapsto \mathcal{B}_{sa}(\mathbb{H})$  is a  $k$  times Fréchet differentiable function, its  $k$ -th derivative  $D^k g(A)$ ,  $A \in \mathcal{B}_{sa}(\mathbb{H})$  can be viewed as a symmetric multilinear operator valued form

$$D^k g(A)(H_1, \dots, H_k) = D^k g(A; H_1, \dots, H_k), H_1, \dots, H_k \in \mathcal{B}_{sa}(\mathbb{H}).$$

- For a  $k$ -linear form  $M : \mathcal{B}_{sa}(\mathbb{H}) \times \dots \times \mathcal{B}_{sa}(\mathbb{H}) \mapsto \mathcal{B}_{sa}(\mathbb{H})$ , define its operator norm as

$$\|M\| := \sup_{\|H_1\|, \dots, \|H_k\| \leq 1} \|M(H_1, \dots, H_k)\|.$$

- The derivatives  $D^k g(A)$  are defined iteratively:

$$D^k g(A)(H_1, \dots, H_{k-1}, H_k) = D(D^{k-1} g(A)(H_1, \dots, H_{k-1}))(H_k).$$



## Lemma

Let  $f \in \mathcal{E}_\sigma \cap L_\infty(\mathbb{R})$ . Then

$$\|D^k f(A)\| \leq \sigma^k \|f\|_{L_\infty(\mathbb{R})}, A \in \mathcal{B}_{sa}(\mathbb{H}),$$

$$\begin{aligned} & \|D^k f(A + H; H_1, \dots, H_k) - D^k f(A; H_1, \dots, H_k)\| \\ & \leq \sigma^{k+1} \|f\|_{L_\infty(\mathbb{R})} \|H_1\| \dots \|H_k\| \|H\| \end{aligned}$$

and

$$\|S_{D^k f(\cdot; H_1, \dots, H_k)}(A; H)\| \leq \frac{\sigma^{k+2}}{2} \|f\|_{L_\infty(\mathbb{R})} \|H_1\| \dots \|H_k\| \|H\|^2.$$

## Lemma

Suppose  $f \in B_{\infty,1}^k(\mathbb{R})$ . Then the function  $\mathcal{B}_{sa}(\mathbb{H}) \ni A \mapsto f(A) \in \mathcal{B}_{sa}(\mathbb{H})$  is  $k$  times Fréchet differentiable and

$$\|D^j f(A)\| \leq 2^j \|f\|_{B_{\infty,1}^j}, A \in \mathcal{B}_{sa}(\mathbb{H}), j = 1, \dots, k.$$

Moreover, if for some  $s \in (k, k + 1]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ , then

$$\|D^k f(A + H) - D^k f(A)\| \leq 2^{k+1} \|f\|_{B_{\infty,1}^s} \|H\|^{s-k}, A, H \in \mathcal{B}_{sa}(\mathbb{H}).$$

# Higher Order Operator Differentiability

For an open set  $G \subset \mathcal{B}_{sa}(\mathbb{H})$ , a  $k$ -times Fréchet differentiable functions  $g : G \mapsto \mathcal{B}_{sa}(\mathbb{H})$  and, for  $s = k + \beta$ ,  $\beta \in (0, 1]$ , define

$$\|g\|_{C^s(G)} := \max_{0 \leq j \leq k} \sup_{A \in G} \|D^j g(A)\| \bigvee \sup_{A, A+H \in G, H \neq 0} \frac{\|D^k g(A+H) - D^k g(A)\|}{\|H\|^\beta}.$$

## Corollary

Suppose that, for some  $s > 0$ ,  $s \in (k, k+1]$ , we have  $f \in B_{\infty,1}^s(\mathbb{R})$ .  
Then

$$\|f\|_{C^s(\mathcal{B}_{sa}(\mathbb{H}))} \leq 2^{k+1} \|f\|_{B_{\infty,1}^s}.$$

# Normal Approximation for Smooth Functions of Sample Covariance

- Let  $\Sigma := \sum_{\lambda \in \sigma(\Sigma)} \lambda P_\lambda$  be the spectral decomposition of  $\Sigma$ ,  $\sigma(\Sigma)$  being the spectrum of  $\Sigma$  and  $P_\lambda$  being the spectral projection corresponding to the eigenvalue  $\lambda$
- $f \in B_{\infty,1}^1(\mathbb{R})$
- $\|B\|_1 < \infty$
- $\sigma_f(\Sigma; B) := \sqrt{2} \|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_2$

# Perturbation Theory: Application to Functions of Sample Covariance (Delta Method)

- $$\langle f(\hat{\Sigma}) - f(\Sigma), B \rangle = \langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle + \langle S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$$
- The linear term  $\langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$  is of the order  $O(n^{-1/2})$  and  $n^{1/2}\langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$  is close in distribution to  $N(0; \sigma_f^2(\Sigma; B))$ .
- For  $s \in (1, 2]$ ,  $\|S_f(\Sigma; \hat{\Sigma} - \Sigma)\| \lesssim \|f\|_{B_{\infty,1}^s} \|\hat{\Sigma} - \Sigma\|^s$ , implying that

$$\begin{aligned} |\langle S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle| &\leq \|B\|_1 \|S_f(\Sigma; \hat{\Sigma} - \Sigma)\| \\ &= O\left(\left(\frac{\mathbf{r}(\Sigma)}{n}\right)^{s/2}\right) = o(n^{-1/2}), \end{aligned}$$

$$\begin{aligned} |\langle \mathbb{E}f(\hat{\Sigma}) - f(\Sigma), B \rangle| &= |\langle \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle| \\ &= O\left(\left(\frac{\mathbf{r}(\Sigma)}{n}\right)^{s/2}\right) = o(n^{-1/2}) \end{aligned}$$

provided that  $\mathbf{r}(\Sigma) = o(n^{1/s-1})$ .

# Perturbation Theory: Application to Functions of Sample Covariance (Delta Method)

- The bounds are sharp, for instance, for  $f(x) = x^2$ ,  $B = u \otimes u$ ,  $s = 2$ :

$$\begin{aligned} \sup_{\|u\| \leq 1} |\langle \mathbb{E}f(\hat{\Sigma}) - f(\Sigma), u \otimes u \rangle| &= \sup_{\|u\| \leq 1} |\langle \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), u \otimes u \rangle| = \\ &= \frac{\|\text{tr}(\Sigma)\Sigma + \Sigma^2\|}{n} \asymp \|\Sigma\|^2 \frac{\mathbf{r}(\Sigma)}{n} \end{aligned}$$

- For  $s = 2$ , the Delta Method works if  $\mathbf{r}(\Sigma) = o(n^{1/2})$ . What if  $\mathbf{r}(\Sigma) \geq n^{1/2}$ ,  $\mathbf{r}(\Sigma) = o(n)$ ?

# Normal Approximation for Smooth Functions of Sample Covariance

Let  $\mathcal{G}(r; a) = \left\{ \Sigma : \|\Sigma\| \leq a, \mathbf{r}(\Sigma) \leq r \right\}$ .

## Theorem

Let  $f \in B_{\infty,1}^s(\mathbb{R})$  for some  $s \in (1, 2]$  and let  $B$  be a linear operator with  $\|B\|_1 < \infty$ . Suppose  $a > 0, \sigma_0^2 > 0$  and

$$r_n = o(n) \text{ as } n \rightarrow \infty.$$

Then

$$\sup_{\Sigma \in \mathcal{G}(r_n; a), \sigma_f(\Sigma; B) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_{\Sigma} \left\{ \frac{n^{1/2} \langle f(\hat{\Sigma}) - \mathbb{E}_{\Sigma} f(\hat{\Sigma}), B \rangle}{\sigma_f(\Sigma, B)} \leq x \right\} - \Phi(x) \right| \rightarrow 0.$$

# Normal Approximation Bounds for Smooth Functions of Sample Covariance

- Let  $\Sigma := \sum_{\lambda \in \sigma(\Sigma)} \lambda P_\lambda$  be the spectral decomposition of  $\Sigma$
- $f \in B_{\infty,1}^1(\mathbb{R})$
- $\|B\|_1 < \infty$
- $\sigma_f(\Sigma; B) := \sqrt{2} \|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_2$
- $\mu_f(\Sigma; B) := \|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_3$
- $\gamma_s(f; \Sigma) := \frac{\|f\|_{B_{\infty,1}^s} \|B\|_1 \|\Sigma\|^s}{\sigma_f(\Sigma; B)}$
- $t_n(\Sigma) := t_{n,s}(f; \Sigma) := \left[ -\log \gamma_s(f; \Sigma) + \frac{s-1}{2} \log \left( \frac{n}{r(\Sigma)} \right) \right] \vee 1.$



# Normal Approximation Bounds for Smooth Functions of Sample Covariance

## Theorem

Let  $f \in B_{\infty,1}^s(\mathbb{R})$  for some  $s \in (1, 2]$ . Then

$$\sup_{x \in \mathbb{R}} \left| \mathbb{P} \left\{ \frac{n^{1/2} \langle f(\hat{\Sigma}) - \mathbb{E}f(\hat{\Sigma}), B \rangle}{\sigma_f(\Sigma, B)} \leq x \right\} - \Phi(x) \right| \lesssim_s \left( \frac{\mu_f(\Sigma; B)}{\sigma_f(\Sigma; B)} \right)^3 \frac{1}{\sqrt{n}}$$
$$+ \gamma_s(f; \Sigma) \left( \left( \frac{\mathbf{r}(\Sigma)}{n} \right)^{(s-1)/2} \vee \left( \frac{t_n(\Sigma)}{n} \right)^{(s-1)/2} \vee \left( \frac{t_n(\Sigma)}{n} \right)^{s-1/2} \right) \sqrt{t_n(\Sigma)}.$$

# Perturbation Theory for Functions of Sample Covariance



$$\langle f(\hat{\Sigma}) - f(\Sigma), B \rangle = \langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle + \langle S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$$

implies that

$$\begin{aligned} & \langle f(\hat{\Sigma}) - \mathbb{E}f(\hat{\Sigma}), B \rangle = \\ & = \langle Df(\Sigma)(\hat{\Sigma} - \Sigma), B \rangle + \langle S_f(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle \\ & = \langle Df(\Sigma)(B), \hat{\Sigma} - \Sigma \rangle + \langle S_f(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle \end{aligned}$$

# Perturbation Theory for Functions of Sample Covariance

- The linear term

$$\begin{aligned} & \langle Df(\Sigma)B, \hat{\Sigma} - \Sigma \rangle \\ &= n^{-1} \sum_{j=1}^n \langle Df(\Sigma; B)X_j, X_j \rangle - \mathbb{E} \langle Df(\Sigma, B)X, X \rangle \end{aligned}$$

is of the order  $O(n^{-1/2})$  and it is approximated by a normal distribution using **Berry-Esseen bound**.

- The centered remainder

$$\langle S_f(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$$

is of the order  $o(n^{-1/2})$  when  $\mathbf{r}(\Sigma) = o(n)$  and it is controlled using Gaussian concentration inequalities.

# Normal Approximation for the Linear Term

- Let  $A \in \mathcal{B}_{sa}(\mathbb{H})$ ,  $\|A\|_1 < \infty$ .
- Denote by  $\lambda_j, j \geq 1$  the eigenvalues of  $\Sigma^{1/2}A\Sigma^{1/2}$
- Then

$$\langle AX, X \rangle \stackrel{d}{=} \sum_{k \geq 1} \lambda_k Z_k^2,$$

where  $Z_1, Z_2, \dots$  are i.i.d. standard normal random variables.

# Normal Approximation for the Linear Term

## Lemma

The following bound holds:

$$\begin{aligned} & \sup_{x \in \mathbb{R}} \left| \mathbb{P} \left\{ \frac{n^{1/2} \langle Df(\Sigma; B), \hat{\Sigma} - \Sigma \rangle}{\sqrt{2} \|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_2} \leq x \right\} - \Phi(x) \right| \\ & \lesssim \left( \frac{\|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_3}{\|\Sigma^{1/2} Df(\Sigma; B) \Sigma^{1/2}\|_2} \right)^3 \frac{1}{\sqrt{n}}. \end{aligned}$$

## Proof.

$$\frac{n^{1/2} \langle Df(\Sigma; B), \hat{\Sigma} - \Sigma \rangle}{\sqrt{2} \|\Sigma^{1/2} Df(\Sigma) \Sigma^{1/2}\|_2} \stackrel{d}{=} \frac{\sum_{j=1}^n \sum_{k \geq 1} \lambda_k (Z_{k,j}^2 - 1)}{\text{Var}^{1/2} \left( \sum_{j=1}^n \sum_{k \geq 1} \lambda_k (Z_{k,j}^2 - 1) \right)},$$

where  $\{Z_{k,j}\}$  are i.i.d. standard normal random variables and  $\lambda_k$  the eigenvalues of  $Df(\Sigma; B)$ . It remains to use Berry-Esseen bound.  $\square$

## Theorem

Suppose that, for some  $s \in (1, 2]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$  and also that  $\mathbf{r}(\Sigma) \lesssim n$ . Then, there exists a constant  $C = C_s > 0$  such that, for all  $t \geq 1$ , with probability at least  $1 - e^{-t}$

$$\left| \left\langle S_f(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E} S_f(\Sigma; \hat{\Sigma} - \Sigma), B \right\rangle \right| \\ \leq C \|f\|_{B_{\infty,1}^s} \|B\|_1 \|\Sigma\|^s \left( \left( \frac{\mathbf{r}(\Sigma)}{n} \right)^{(s-1)/2} V \left( \frac{t}{n} \right)^{(s-1)/2} V \left( \frac{t}{n} \right)^{s-1/2} \right) \sqrt{\frac{t}{n}}$$

**Note:** the centered remainder is  $o_{\mathbb{P}}(n^{-1/2})$  provided that  $\mathbf{r}(\Sigma) = o(n)$ .

# Concentration of the Remainder

- $g : \mathcal{B}_{sa}(\mathbb{H}) \mapsto \mathbb{R}$  Fréchet differentiable function with respect to the operator norm with derivative  $Dg(A; H)$ ,  $H \in \mathcal{B}_{sa}(\mathbb{H})$ .
- $S_g(A; H)$  the remainder of the first order Taylor expansion of  $g$  :

$$S_g(A; H) := g(A + H) - g(A) - Dg(A; H), A, H \in \mathcal{B}_{sa}(\mathbb{H}).$$

- Let  $s \in [1, 2]$ . Assume there exists a constant  $L_{g,s} > 0$  such that, for all  $\Sigma \in \mathcal{C}_+(\mathbb{H})$ ,  $H, H' \in \mathcal{B}_{sa}(\mathbb{H})$ ,

$$|S_g(\Sigma; H') - S_g(\Sigma; H)| \leq L_{g,s}(\|H\| \vee \|H'\|)^{s-1} \|H' - H\|.$$

## Theorem

*There exists a constant  $K_s > 0$  such that for all  $t \geq 1$  with probability at least  $1 - e^{-t}$*

$$\begin{aligned} & |S_g(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_g(\Sigma; \hat{\Sigma} - \Sigma)| \\ & \leq K_s L_{g,s} \|\Sigma\|^s \left( \left( \frac{\mathbf{r}(\Sigma)}{n} \right)^{(s-1)/2} V \left( \frac{\mathbf{r}(\Sigma)}{n} \right)^{s-1/2} V \left( \frac{t}{n} \right)^{(s-1)/2} V \left( \frac{t}{n} \right)^{s-1/2} \right) \end{aligned}$$



# Sketch of the Proof

- $\varphi : \mathbb{R} \mapsto \mathbb{R}$ ,  $\varphi(u) = 1$ ,  $u \leq 1$ ,  $\varphi(u) = 0$ ,  $u \geq 2$ ,  
 $\varphi(u) = 2 - u$ ,  $u \in (1, 2)$
- $E := \hat{\Sigma} - \Sigma$
- For  $\delta > 0$ , define

$$h(X_1, \dots, X_n) := S_g(\Sigma; E) \varphi\left(\frac{\|E\|}{\delta}\right).$$

## Lemma

The following bound holds with some constant  $C_s > 0$  for all  $X_1, \dots, X_n, X'_1, \dots, X'_n \in \mathbb{H}$  :

$$\begin{aligned} & |h(X_1, \dots, X_n) - h(X'_1, \dots, X'_n)| \\ & \leq \frac{C_s L_{g,s} (\|\Sigma\|^{1/2} + \sqrt{\delta}) \delta^{s-1}}{\sqrt{n}} \left( \sum_{j=1}^n \|X_j - X'_j\|^2 \right)^{1/2}. \end{aligned}$$



$$\begin{aligned} & |h(X_1, \dots, X_n) - h(X'_1, \dots, X'_n)| \\ & \leq |\mathcal{S}_g(\Sigma, E) - \mathcal{S}_g(\Sigma, E')| + \frac{1}{\delta} |\mathcal{S}_g(\Sigma, E')| \|E - E'\| \\ & \leq L_{g,s} (\|E\| \vee \|E'\|)^{s-1} \|E' - E\| + L_{g,s} \frac{1}{\delta} \|E'\|^s \|E' - E\|. \end{aligned}$$

- If both  $\|E\| \leq 2\delta$  and  $\|E'\| \leq 2\delta$ , then

$$|h(X_1, \dots, X_n) - h(X'_1, \dots, X'_n)| \leq (2^{s-1} + 2^s) L_{g,s} \delta^{s-1} \|E' - E\|$$

with similar bounds holding in other cases.

# Sketch of the Proof

$$\begin{aligned}\|E' - E\| &= \left\| n^{-1} \sum_{j=1}^n X_j \otimes X_j - n^{-1} \sum_{j=1}^n X'_j \otimes X'_j \right\| \\ &\leq \left\| n^{-1} \sum_{j=1}^n (X_j - X'_j) \otimes X_j \right\| + \left\| n^{-1} \sum_{j=1}^n X'_j \otimes (X_j - X'_j) \right\| \\ &= \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X_j - X'_j, u \rangle \langle X_j, v \rangle \right| + \sup_{\|u\|, \|v\| \leq 1} \left| n^{-1} \sum_{j=1}^n \langle X'_j, u \rangle \langle X_j - X'_j, v \rangle \right|\end{aligned}$$

# Sketch of the Proof

$$\begin{aligned} &\leq \sup_{\|u\| \leq 1} \left( n^{-1} \sum_{j=1}^n \langle X_j - X'_j, u \rangle^2 \right)^{1/2} \sup_{\|v\| \leq 1} \left( n^{-1} \sum_{j=1}^n \langle X_j, v \rangle^2 \right)^{1/2} \\ &+ \sup_{\|u\| \leq 1} \left( n^{-1} \sum_{j=1}^n \langle X'_j, u \rangle^2 \right)^{1/2} \sup_{\|v\| \leq 1} \left( n^{-1} \sum_{j=1}^n \langle X_j - X'_j, v \rangle^2 \right)^{1/2} \\ &\leq \frac{\|\hat{\Sigma}\|^{1/2} + \|\hat{\Sigma}'\|^{1/2}}{\sqrt{n}} \left( \sum_{j=1}^n \|X_j - X'_j\|^2 \right)^{1/2} \\ &\leq (2\|\Sigma\|^{1/2} + \|E\|^{1/2} + \|E'\|^{1/2})\Delta, \end{aligned}$$

where  $\Delta := \frac{1}{\sqrt{n}} \left( \sum_{j=1}^n \|X_j - X'_j\|^2 \right)^{1/2}$ .

By a simple further algebra,

$$\|E' - E\| \wedge \delta \leq 4\|\Sigma\|^{1/2}\Delta \sqrt{(4\sqrt{2} + 2)\sqrt{\delta}\Delta},$$

which together with the bound

$$|h(X_1, \dots, X_n) - h(X'_1, \dots, X'_n)| \lesssim_s L_{g,s} \delta^{s-1} \|E' - E\|$$

implies the claim of the lemma.

# Sketch of the Proof

- For a given  $t > 0$ , let

$$\delta = \delta_n(t) := \mathbb{E}\|\hat{\Sigma} - \Sigma\| + C\|\Sigma\| \left[ \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee 1 \right) \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right]$$

with constant  $C > 0$  such that  $\mathbb{P}\{\|\hat{\Sigma} - \Sigma\| \geq \delta_n(t)\} \leq e^{-t}$ .

- Let  $M := \text{Med}(S_g(\Sigma; \hat{\Sigma} - \Sigma))$
- By Gaussian concentration inequality, probability at least  $1 - e^{-t}$

$$|h(X_1, \dots, X_n) - M| \lesssim_s L_{g,s} \delta^{s-1} (\|\Sigma\|^{1/2} + \delta^{1/2}) \|\Sigma\|^{1/2} \sqrt{\frac{t}{n}}.$$

- On the event  $\{\|\hat{\Sigma} - \Sigma\| \leq \delta\}$ , replace  $h(X_1, \dots, X_n)$  by  $S_g(\Sigma; \hat{\Sigma} - \Sigma)$

# Assumptions on the Loss Function

Let  $\mathcal{L}$  be the class of functions  $\ell : \mathbb{R} \mapsto \mathbb{R}_+$  such that

- $\ell(0) = 0$
- $\ell(u) = \ell(-u), u \in \mathbb{R}$
- $\ell$  is nondecreasing and convex on  $\mathbb{R}_+$
- For some constants  $c_1, c_2 > 0$

$$\ell(u) \leq c_1 e^{c_2 u}, u \geq 0.$$



# When is Plug-In Estimator Asymptotically Efficient?

## Theorem

Suppose, for some  $s \in (1, 2]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ . Let  $B$  be a nuclear operator and let  $a > 0, \sigma_0 > 0$ . Suppose that  $r_n > 1$  and  $r_n = o(n^{1-\frac{1}{s}})$  as  $n \rightarrow \infty$ . Then

$$\sup_{\Sigma \in \mathcal{G}(r_n; a), \sigma_f(\Sigma; B) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_{\Sigma} \left\{ \frac{n^{1/2} (\langle f(\hat{\Sigma}), B \rangle - \langle f(\Sigma), B \rangle)}{\sigma_f(\Sigma; B)} \leq x \right\} - \Phi(x) \right| \rightarrow 0$$

as  $n \rightarrow \infty$ . Moreover, under the same assumptions on  $f$  and  $r_n$ , and for any loss function  $\ell \in \mathcal{L}$

$$\sup_{\Sigma \in \mathcal{G}(r_n; a), \sigma_f(\Sigma; B) \geq \sigma_0} \left| \mathbb{E}_{\Sigma} \ell \left( \frac{n^{1/2} (\langle f(\hat{\Sigma}), B \rangle - \langle f(\Sigma), B \rangle)}{\sigma_f(\Sigma; B)} \right) - \mathbb{E} \ell(Z) \right| \rightarrow 0$$

as  $n \rightarrow \infty$ , where  $Z$  is a standard normal random variable.

# Efficient Estimation of $\langle f(\Sigma), B \rangle$ : A Lower Bound

## Theorem

Suppose  $f \in B_{\infty,1}^1(\mathbb{R})$ . Suppose  $r_n > 1, a > 1, \sigma_0 > 0$  are such that, for some  $1 < a' < a$  and  $\sigma'_0 > \sigma_0$  and for all large enough  $n$ ,

$$\mathcal{G}(r_n, a') \cap \left\{ \Sigma : \sigma_f(\Sigma; B) \geq \sigma'_0 \right\} \neq \emptyset.$$

Then the following bound holds:

$$\liminf_n \inf_{T_n} \sup_{\Sigma \in \mathcal{G}(r_n; a), \sigma_f(\Sigma; B) \geq \sigma_0} \frac{n \mathbb{E}_{\Sigma} \left( T_n(X_1, \dots, X_n) - \langle f(\Sigma), B \rangle \right)^2}{\sigma_f^2(\Sigma; B)} \geq 1,$$

where the infimum is taken over all sequences of estimators  $T_n = T_n(X_1, \dots, X_n)$ .

## Part 3.

# Wishart Operators, Bootstrap Chains, Invariant Functions and Bias Reduction

- Our next goal is to find an estimator  $g(\hat{\Sigma})$  of  $f(\Sigma)$  with a small bias  $\mathbb{E}_{\Sigma}g(\hat{\Sigma}) - f(\Sigma)$  (of the order  $o(n^{-1/2})$ ) and such that

$$n^{1/2}(\langle g(\hat{\Sigma}), B \rangle - \langle \mathbb{E}_{\Sigma}g(\hat{\Sigma}), B \rangle)$$

is asymptotically normal.

- To this end, one has to find a sufficiently smooth approximate solution of the equation

$$\mathbb{E}_{\Sigma}g(\hat{\Sigma}) = f(\Sigma), \Sigma \in \mathcal{C}_+^d.$$

# Wishart Operators

- $Tg(\Sigma) := \mathbb{E}_{\Sigma}g(\hat{\Sigma}) = \int_{\mathcal{C}_+^d} g(V)P(\Sigma; dV), \Sigma \in \mathcal{C}_+^d,$   
where Markov kernel  $P(\Sigma; \cdot)$  is a rescaled Wishart distribution  $\mathcal{W}_d(\Sigma; n)$ :

$$P(\Sigma; A) := \mathbb{P}_{\Sigma}\{\hat{\Sigma} \in A\}, A \subset \mathcal{C}_+^d.$$

- For  $d \leq n$ ,  $P(\Sigma; dV) = np(\Sigma; nV)dV,$

$$p(\Sigma; V) := \frac{1}{2^{nd/2}(\det(\Sigma))^{n/2}\Gamma_d\left(\frac{n}{2}\right)} (\det(V))^{(n-d-1)/2} \exp\left\{-\frac{1}{2}\text{tr}(\Sigma^{-1}V)\right\},$$

where  $\Gamma_d$  is the multivariate gamma function:

$$\Gamma_d\left(\frac{n}{2}\right) := \pi^{d(d-1)/4} \prod_{j=1}^d \Gamma\left(\frac{n}{2} - \frac{j-1}{2}\right).$$

# Bias Reduction

- To find an estimator of  $f(\Sigma)$  with a small bias, one needs to solve (approximately) the following integral equation
- the **Wishart equation**:

$$\mathcal{T}g(\Sigma) = f(\Sigma), \Sigma \in \mathcal{C}_+^d.$$

- Bias operator:  $\mathcal{B} := \mathcal{T} - \mathcal{I}$ .
- Formally, the solution of the Wishart equation is given by Neumann series:

$$g(\Sigma) = (\mathcal{I} + \mathcal{B})^{-1}f(\Sigma) = (\mathcal{I} - \mathcal{B} + \mathcal{B}^2 - \dots)f(\Sigma)$$

- Given a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$ , define

$$f_k(\Sigma) := \sum_{j=0}^k (-1)^j \mathcal{B}^j f(\Sigma) := f(\Sigma) + \sum_{j=1}^k (-1)^j \mathcal{B}^j f(\Sigma)$$

## Proposition

The bias of estimator  $f_k(\hat{\Sigma})$  of  $f(\Sigma)$  is given by the following formula:

$$\mathbb{E}_{\Sigma} f_k(\hat{\Sigma}) - f(\Sigma) = (-1)^k \mathcal{B}^{k+1} f(\Sigma).$$

## Proof.

$$\begin{aligned} \mathbb{E}_{\Sigma} f_k(\hat{\Sigma}) - g(\Sigma) &= \mathcal{T} f_k(\Sigma) - f(\Sigma) = (\mathcal{I} + \mathcal{B}) f_k(\Sigma) - f(\Sigma) \\ &= \sum_{j=0}^k (-1)^j \mathcal{B}^j f(\Sigma) - \sum_{j=1}^{k+1} (-1)^j \mathcal{B}^j f(\Sigma) - f(\Sigma) = (-1)^k \mathcal{B}^{k+1} f(\Sigma). \end{aligned}$$



# Bootstrap Chain

- $\mathcal{T}g(\Sigma) := \mathbb{E}_{\Sigma}g(\hat{\Sigma}) = \int_{\mathcal{C}_+^d} g(V)P(\Sigma; dV), \Sigma \in \mathcal{C}_+^d$
- $\mathcal{T}^k g(\Sigma) = \mathbb{E}_{\Sigma}g(\hat{\Sigma}^{(k)}), \Sigma \in \mathcal{C}_+^d$ , where

$$\hat{\Sigma}^{(0)} = \Sigma \rightarrow \hat{\Sigma}^{(1)} = \hat{\Sigma} \rightarrow \hat{\Sigma}^{(2)} \rightarrow \dots$$

is a Markov chain in  $\mathcal{C}_+^d$  with transition probability kernel  $P(\cdot; \cdot)$ .

- Note that  $\hat{\Sigma}^{(j+1)}$  is the sample covariance based on  $n$  i.i.d. observations  $\sim N(0; \hat{\Sigma}^{(j)})$  (conditionally on  $\hat{\Sigma}^{(j)}$ )
- Conditionally on  $\hat{\Sigma}^{(j)}$ , with a “high probability”,

$$\|\hat{\Sigma}^{(j+1)} - \hat{\Sigma}^{(j)}\| \lesssim \|\hat{\Sigma}^{(j)}\| \sqrt{\frac{d}{n}}$$

- The Markov Chain  $\{\hat{\Sigma}^{(j)}, j = 0, 1, \dots\}$  will be called *the Bootstrap Chain*.



# $k$ -th order difference

- $k$ -th order difference along the Markov chain: by Newton binomial formula,

$$\begin{aligned} \mathcal{B}^k f(\Sigma) &= (\mathcal{T} - \mathcal{I})^k f(\Sigma) = \sum_{j=0}^k (-1)^{k-j} \binom{k}{j} \mathcal{T}^j f(\Sigma) \\ &= \mathbb{E}_{\Sigma} \sum_{j=0}^k (-1)^{k-j} \binom{k}{j} f(\hat{\Sigma}^{(j)}) \end{aligned}$$

- $\sum_{j=0}^k (-1)^{k-j} \binom{k}{j} f(\hat{\Sigma}^{(j)})$  is the  $k$ -th order difference of  $f$  along the trajectory of the Bootstrap Chain.
- **Question.** Suppose  $f$  is of smoothness  $k$ . Is  $\mathcal{B}^k f(\Sigma)$  of the order  $\left(\sqrt{\frac{d}{n}}\right)^k$ ?

# Orthogonally Invariant Functions

- A function  $g \in L_\infty(\mathcal{C}_+^d)$  is called *orthogonally invariant* iff, for all orthogonal transformations  $U$  of  $\mathbb{R}^d$ ,

$$g(U\Sigma U^{-1}) = g(\Sigma), \Sigma \in \mathcal{C}_+^d.$$

- Any orthogonally invariant function  $g$  could be represented as  $g(\Sigma) = \varphi(\lambda_1(\Sigma), \dots, \lambda_d(\Sigma))$ , where  $\lambda_1(\Sigma) \geq \dots \geq \lambda_d(\Sigma)$  are the eigenvalues of  $\Sigma$  and  $\varphi$  is a symmetric function of  $d$  variables.
- A typical example:  $g(\Sigma) = \text{tr}(\psi(\Sigma))$  for a function of real variable  $\psi$ .
- Let  $L_\infty^O(\mathcal{C}_+^d)$  be the space of all orthogonally invariant functions from  $L_\infty(\mathcal{C}_+^d)$ .

# Orthogonally Invariant Functions

## Proposition

If  $g \in L_{\infty}^O(\mathcal{C}_+^d)$ , then  $\mathcal{T}g \in L_{\infty}^O(\mathcal{C}_+^d)$  and  $\mathcal{B}g \in L_{\infty}^O(\mathcal{C}_+^d)$ .

## Proof.

Indeed, the transformation  $\Sigma \mapsto U\Sigma U^{-1}$  is a bijection of  $\mathcal{C}_+^d$ ,

$$\mathcal{T}g(U\Sigma U^{-1}) = \mathbb{E}_{U\Sigma U^{-1}} g(\hat{\Sigma}) = \mathbb{E}_{\Sigma} g(U\hat{\Sigma}U^{-1}) = \mathbb{E}_{\Sigma} g(\hat{\Sigma}) = \mathcal{T}g(\Sigma)$$

and the function  $\mathcal{T}g$  is uniformly bounded. □ □

# Orthogonally Equivariant Functions

$g : \mathcal{C}_+^d \mapsto \mathcal{B}_{sa}(\mathbb{R}^d)$  is called *orthogonally equivariant* iff for all orthogonal transformations  $U$

$$g(U\Sigma U^{-1}) = Ug(\Sigma)U^{-1}, \Sigma \in \mathcal{C}_+^d.$$

$g : \mathcal{C}_+^d \mapsto \mathcal{B}_{sa}(\mathbb{R}^d)$  is continuously differentiable in  $\mathcal{C}_+^d$  iff there exists a uniformly bounded, Lipschitz with respect to the operator norm and continuously differentiable extension of  $g$  to an open set  $G$ ,

$$\mathcal{C}_+^d \subset G \subset \mathcal{B}_{sa}(\mathbb{R}^d).$$

## Proposition

If  $g : \mathcal{C}_+^d \mapsto \mathbb{R}$  is orthogonally invariant and continuously differentiable in  $\mathcal{C}_+^d$  with derivative  $Dg$ , then  $Dg$  is orthogonally equivariant.

# Orthogonally Equivariant Functions

## Proof.

First suppose that  $\Sigma$  is positively definite (then extend to  $\mathcal{C}_+^d$  by continuity). For all  $H \in \mathcal{B}_{sa}(\mathbb{R}^d)$ ,

$$\begin{aligned}\langle Dg(U\Sigma U^{-1}), H \rangle &= \lim_{t \rightarrow 0} \frac{g(U\Sigma U^{-1} + tH) - g(U\Sigma U^{-1})}{t} \\ &= \lim_{t \rightarrow 0} \frac{g(U(\Sigma + tU^{-1}HU)U^{-1}) - g(U\Sigma U^{-1})}{t} \\ &= \lim_{t \rightarrow 0} \frac{g(\Sigma + tU^{-1}HU) - g(\Sigma)}{t} \\ &= \langle Dg(\Sigma), U^{-1}HU \rangle = \langle UDg(\Sigma)U^{-1}, H \rangle.\end{aligned}$$



# “Lifting” Operator $\mathcal{D}$ and Reduction to Orthogonally Invariant Functions

- Define the following differential operator (“lifting” operator):

$$\mathcal{D}g(\Sigma) := \Sigma^{1/2} Dg(\Sigma) \Sigma^{1/2}$$

acting on continuously differentiable functions in  $\mathcal{C}_+^d$ .

- Suppose  $f(x) = x\psi'(x)$
- Let  $g(\Sigma) := \text{tr}(\psi(\Sigma))$
- $g$  is orthogonally invariant function on  $\mathcal{C}_+^d$
- $Dg(\Sigma) = \psi'(\Sigma)$
- $\mathcal{D}g(\Sigma) = f(\Sigma)$

# "Lifting" Operator $\mathcal{D}$ and Reduction to Invariant Functions

- Suppose, for some  $\delta > 0$ ,  $\sigma(\Sigma) \subset [2\delta, \infty)$ .
- Let  $\gamma_\delta(x) = \gamma(x/\delta)$ ,  $\gamma : \mathbb{R} \mapsto [0, 1]$  be a nondecreasing  $C^\infty$  function,  $\gamma(x) = 0, x \leq 1/2$ ,  $\gamma(x) = 1, x \geq 1$ .
- Define  $f_\delta(x) = f(x)\gamma_\delta(x)$ ,  $x \in \mathbb{R}$ .
- Then,  $f(\Sigma) = f_\delta(\Sigma)$  and, for all  $\Sigma$  with  $\sigma(\Sigma) \subset [2\delta, \infty)$ ,  $Df(\Sigma) = Df_\delta(\Sigma)$ .
- Let  $\varphi(x) := \int_0^x \frac{f_\delta(t)}{t} dt$ ,  $x \geq 0$ ,  $\varphi(x) = 0, x < 0$ .
- Clearly,  $f_\delta(x) = x\varphi'(x)$ ,  $x \in \mathbb{R}$ .
- Let  $g(C) := \text{tr}(\varphi(C))$ ,  $C \in \mathcal{B}_{sa}(\mathbb{R}^d)$ . Then

$$Dg(C) = C^{1/2}\varphi'(C)C^{1/2} = f_\delta(C), C \in \mathcal{C}_+^d.$$

- Moreover,

$$\|Dg\|_{C^s} \lesssim 2^{k+1}(\delta^{-1-s} \vee \delta^{-1})\|f\|_{B_{\infty,1}^s}.$$

# Operator $\mathcal{D}$ and its Commutativity Properties

## Proposition

Suppose  $d \lesssim n$ . For all functions  $g \in L_{\infty}^O(\mathcal{C}_+^d)$  that are continuously differentiable in  $\mathcal{C}_+^d$  with a uniformly bounded derivative  $Dg$  and for all  $\Sigma \in \mathcal{C}_+^d$

$$\mathcal{D}Tg(\Sigma) = T\mathcal{D}g(\Sigma) \text{ and } \mathcal{D}B_g(\Sigma) = B\mathcal{D}g(\Sigma).$$



# Operator $\mathcal{D}$ and its Commutativity Properties (Proof)

- Note that  $\hat{\Sigma} \stackrel{d}{=} \Sigma^{1/2} W \Sigma^{1/2}$ , where  $W$  is the sample covariance based on i.i.d. standard normal random variables  $Z_1, \dots, Z_n$  in  $\mathbb{R}^d$ .
- Let  $\Sigma^{1/2} W^{1/2} = RU$  be the polar decomposition of  $\Sigma^{1/2} W^{1/2}$  with positively semidefinite  $R$  and orthogonal  $U$ .
- Then,

$$\hat{\Sigma} = \Sigma^{1/2} W \Sigma^{1/2} = \Sigma^{1/2} W^{1/2} W^{1/2} \Sigma^{1/2} = RUU^{-1}R = R^2$$

and

$$\begin{aligned} W^{1/2} \Sigma W^{1/2} &= W^{1/2} \Sigma^{1/2} \Sigma^{1/2} W^{1/2} \\ &= U^{-1} R R U = U^{-1} R^2 U = U^{-1} \Sigma^{1/2} W \Sigma^{1/2} U = U^{-1} \hat{\Sigma} U. \end{aligned}$$

# Operator $\mathcal{D}$ and its Commutativity Properties (Proof)

- Since  $g$  is orthogonally invariant, we have

$$\mathcal{T}g(\Sigma) = \mathbb{E}_{\Sigma}g(\hat{\Sigma}) = \mathbb{E}g(\Sigma^{1/2}W\Sigma^{1/2}) = \mathbb{E}g(W^{1/2}\Sigma W^{1/2}), \Sigma \in \mathcal{C}_+^d.$$

- For simplicity, assume that  $\Sigma$  is positively definite.
- Let  $H \in \mathcal{B}_{sa}(\mathbb{R}^d)$  and  $\Sigma_t := \Sigma + tH, t > 0$ . Note that

$$\begin{aligned} & D\mathcal{T}g(\Sigma) \\ &= \lim_{t \rightarrow 0} \frac{\mathcal{T}g(\Sigma_t) - \mathcal{T}g(\Sigma)}{t} \\ &= \lim_{t \rightarrow 0} \mathbb{E} \frac{g(W^{1/2}\Sigma_t W^{1/2}) - g(W^{1/2}\Sigma W^{1/2})}{t} \\ &= \mathbb{E} \langle W^{1/2} Dg(W^{1/2}\Sigma W^{1/2}) W^{1/2}, H \rangle \\ &= \langle \mathbb{E} W^{1/2} Dg(W^{1/2}\Sigma W^{1/2}) W^{1/2}, H \rangle. \end{aligned}$$

# Operator $\mathcal{D}$ and its Commutativity Properties (Proof)

- It follows that

$$DTg(\Sigma) = \mathbb{E}W^{1/2}Dg(W^{1/2}\Sigma W^{1/2})W^{1/2}.$$

- Since  $W^{1/2}\Sigma W^{1/2} = U^{-1}\hat{\Sigma}U$  and  $Dg$  is an orthogonally equivariant function, we get

$$Dg(W^{1/2}\Sigma W^{1/2}) = U^{-1}Dg(\hat{\Sigma})U.$$

- Therefore,

$$\begin{aligned}DTg(\Sigma) &= \Sigma^{1/2}DTg(\Sigma)\Sigma^{1/2} \\ &= \Sigma^{1/2}\mathbb{E}(W^{1/2}Dg(W^{1/2}\Sigma W^{1/2})W^{1/2})\Sigma^{1/2} \\ &= \mathbb{E}(\Sigma^{1/2}W^{1/2}Dg(W^{1/2}\Sigma W^{1/2})W^{1/2}\Sigma^{1/2}) \\ &= \mathbb{E}(\Sigma^{1/2}W^{1/2}U^{-1}Dg(\hat{\Sigma})UW^{1/2}\Sigma^{1/2}) \\ &= \mathbb{E}(RUU^{-1}Dg(\hat{\Sigma})UU^{-1}R) = \mathbb{E}(RDg(\hat{\Sigma})R) = \mathbb{E}_{\Sigma}(\hat{\Sigma}^{1/2}Dg(\hat{\Sigma})\hat{\Sigma}^{1/2}) \\ &= \mathbb{E}_{\Sigma}\mathcal{D}g(\hat{\Sigma}) = \mathcal{T}\mathcal{D}g(\Sigma). \quad \square\end{aligned}$$

# Properties of Operators $\mathcal{T}^k$ and $\mathcal{B}^k$

Let  $W_1, \dots, W_k, \dots$  be i.i.d. copies of  $W$ .

## Proposition

Suppose  $d \lesssim n$ . Then, for all  $g \in L_\infty^O(\mathcal{C}_+^d)$  and for all  $k \geq 1$ ,

$$\mathcal{T}^k g(\Sigma) = \mathbb{E} g(W_k^{1/2} \dots W_1^{1/2} \Sigma W_1^{1/2} \dots W_k^{1/2})$$

and

$$\mathcal{B}^k g(\Sigma) = \mathbb{E} \sum_{I \subset \{1, \dots, k\}} (-1)^{k-|I|} g(A_I^* \Sigma A_I),$$

where  $A_I := \prod_{i \in I} W_i^{1/2}$ .

# Properties of Operators $\mathcal{T}^k$ and $\mathcal{B}^k$ (Proof)

- Since  $\hat{\Sigma} \stackrel{d}{=} \Sigma^{1/2} W \Sigma^{1/2}$ ,  $W^{1/2} \Sigma W^{1/2} = U^{-1} \Sigma^{1/2} W \Sigma^{1/2} U$ , where  $U$  is an orthogonal operator, and  $g$  is orthogonally invariant, we have

$$\mathcal{T}g(\Sigma) = \mathbb{E}_{\Sigma} g(\hat{\Sigma}) = \mathbb{E} g(W^{1/2} \Sigma W^{1/2}).$$

- Orthogonal invariance of  $g$  implies the same property of  $\mathcal{T}g$  and, by induction, of  $\mathcal{T}^k g$  for all  $k \geq 1$ . Then, also by induction,

$$\mathcal{T}^k g(\Sigma) = \mathbb{E} g(W_k^{1/2} \dots W_1^{1/2} \Sigma W_1^{1/2} \dots W_k^{1/2}).$$

- If  $I \subset \{1, \dots, k\}$  with  $|I| = \text{card}(I) = j$  and  $A_I = \prod_{i \in I} W_i^{1/2}$ , it implies that

$$\mathcal{T}^j g(\Sigma) = \mathbb{E} g(A_I^* \Sigma A_I).$$

# Properties of Operators $\mathcal{T}^k$ and $\mathcal{B}^k$ (Proof)

- Recall that

$$\mathcal{B}^k g(\Sigma) = (\mathcal{T} - \mathcal{I})^k g(\Sigma) = \sum_{j=0}^k (-1)^{k-j} \binom{k}{j} \mathcal{T}^j g(\Sigma)$$

- Therefore,

$$\mathcal{B}^k g(\Sigma) = \mathbb{E} \sum_{I \subset \{1, \dots, k\}} (-1)^{k-|I|} g(A_I^* \Sigma A_I).$$



## Proposition

Suppose that  $d \lesssim n$  and that  $g$  is continuously differentiable in  $\mathcal{C}_+^d$  with a uniformly bounded derivative  $Dg$ . Then

- 1  $D\mathcal{B}^k g(\Sigma) = \mathbb{E} \sum_{I \subset \{1, \dots, k\}} (-1)^{k-|I|} A_I Dg(A_I^* \Sigma A_I) A_I^*$ .
- 2 For all  $\Sigma \in \mathcal{C}_+^d$ ,

$$D\mathcal{T}^k g(\Sigma) = \mathcal{T}^k Dg(\Sigma) \text{ and } D\mathcal{B}^k g(\Sigma) = \mathcal{B}^k Dg(\Sigma).$$

3

$$\begin{aligned} \mathcal{B}^k Dg(\Sigma) &= D\mathcal{B}^k g(\Sigma) \\ &= \mathbb{E} \left( \sum_{I \subset \{1, \dots, k\}} (-1)^{k-|I|} \Sigma^{1/2} A_I Dg(A_I^* \Sigma A_I) A_I^* \Sigma^{1/2} \right). \end{aligned}$$

# An Integral Representation of Operator $\mathcal{B}^k \mathcal{D}g(\Sigma)$

- Linear interpolation between  $I$  and  $W_1^{1/2}, \dots, W_k^{1/2}$  :

$$V_j(t_j) := I + t_j(W_j^{1/2} - I), t_j \in [0, 1], 1 \leq j \leq k.$$

- For all  $j = 1, \dots, k, t_j \in [0, 1], V_j(t_j) \in \mathcal{C}_+^d$ .



$$R := R(t_1, \dots, t_k) = V_1(t_1) \dots V_k(t_k),$$

$$L := L(t_1, \dots, t_k) = V_k(t_k) \dots V_1(t_1) = R^*,$$

$$S := S(t_1, \dots, t_k) = L(t_1, \dots, t_k) \Sigma R(t_1, \dots, t_k), (t_1, \dots, t_k) \in [0, 1]^k.$$

- Let

$$\varphi(t_1, \dots, t_k) := \Sigma^{1/2} R(t_1, \dots, t_k) \mathcal{D}g(S(t_1, \dots, t_k)) L(t_1, \dots, t_k) \Sigma^{1/2}.$$

- Let  $(t_1, \dots, t_k) \in \{0, 1\}^k, I := \{j : 1 \leq j \leq k, t_j = 1\}$ . Then

$$\varphi(t_1, \dots, t_k) = \Sigma^{1/2} A_I \mathcal{D}g(A_I^* \Sigma A_I) A_I^* \Sigma^{1/2}.$$



# An Integral Representation of Operator $\mathcal{B}^k \mathcal{D}g(\Sigma)$

## Proposition

Suppose  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 1$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 1$ . Then the function  $\varphi$  is  $k$  times continuously differentiable in  $[0, 1]^k$  and

$$\mathcal{B}^k \mathcal{D}g(\Sigma) = \mathbb{E} \int_0^1 \dots \int_0^1 \frac{\partial^k \varphi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} dt_1 \dots dt_k, \Sigma \in \mathcal{C}_+^d.$$

# An Integral Representation of Operator $\mathcal{B}^k \mathcal{D}g(\Sigma)$ (Proof)

- Given a function  $\phi : [0, 1]^k \mapsto \mathbb{R}$ , define for  $1 \leq i \leq k$  finite difference operators

$$\begin{aligned}\Delta_i \phi(t_1, \dots, t_k) \\ := \phi(t_1, \dots, t_{i-1}, 1, t_{i+1}, \dots, t_k) - \phi(t_1, \dots, t_{i-1}, 0, t_{i+1}, \dots, t_k)\end{aligned}$$

- Then

$$\Delta_1 \dots \Delta_k \phi = \sum_{(t_1, \dots, t_k) \in \{0, 1\}^k} (-1)^{k - (t_1 + \dots + t_k)} \phi(t_1, \dots, t_k).$$

- It is well known that if  $\phi$  is  $k$  times continuously differentiable in  $[0, 1]^k$ , then

$$\Delta_1 \dots \Delta_k \phi = \int_0^1 \dots \int_0^1 \frac{\partial^k \phi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} dt_1 \dots dt_k.$$

# An Integral Representation of Operator $\mathcal{B}^k \mathcal{D}g(\Sigma)$ (Proof)



$$\begin{aligned}\mathcal{B}^k \mathcal{D}g(\Sigma) &= \mathcal{D}\mathcal{B}^k g(\Sigma) \\ &= \mathbb{E} \left( \sum_{I \subset \{1, \dots, k\}} (-1)^{k-|I|} \Sigma^{1/2} A_I Dg(A_I^* \Sigma A_I) A_I^* \Sigma^{1/2} \right) \\ &= \sum_{(t_1, \dots, t_k) \in \{0, 1\}^k} (-1)^{k-(t_1 + \dots + t_k)} \varphi(t_1, \dots, t_k) \\ &= \mathbb{E} \Delta_1 \dots \Delta_k \varphi.\end{aligned}$$

- Since  $Dg$  is  $k$  times continuously differentiable and the functions  $S(t_1, \dots, t_k)$ ,  $R(t_1, \dots, t_k)$  are polynomials with respect to  $t_1, \dots, t_k$ , the function  $\varphi$  is  $k$  times continuously differentiable in  $[0, 1]^k$ .
- This implies

$$\mathcal{B}^k \mathcal{D}g(\Sigma) = \mathbb{E} \int_0^1 \dots \int_0^1 \frac{\partial^k \varphi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} dt_1 \dots dt_k, \Sigma \in \mathcal{C}_+^d.$$

# A bound on $\mathcal{B}^k \mathcal{D}g(\Sigma)$

## Theorem

Suppose that  $k \leq d \leq n$  and that  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 1$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 1$ . Then, for some  $C > 1$ ,

$$\|\mathcal{B}^k \mathcal{D}g(\Sigma)\| \leq C^{k^2} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{k/2}.$$

# Bounds on the bias of $\mathcal{D}g_k(\hat{\Sigma})$

## Corollary

Suppose that  $k + 1 \leq d \leq n$  and that  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 2$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 2$ . Then, for some  $C > 1$ ,

$$\begin{aligned} & \|\mathbb{E}_\Sigma \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| \\ & \leq C^{(k+1)^2} \max_{1 \leq j \leq k+2} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+2} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{(k+1)/2}. \end{aligned}$$

If, for some  $\alpha \in (1/2, 1)$ ,  $2 \log n \leq d \leq n^\alpha$  and  $k > \frac{\alpha}{1-\alpha}$ , then

$$\|\mathbb{E}_\Sigma \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| = o(n^{-1/2}).$$

# Further bound on the bias of $\mathcal{D}g_k(\hat{\Sigma})$

## Theorem

Suppose  $g \in L_{\infty}^O(\mathcal{C}_+^d)$  is  $k + 2$  times continuously differentiable for some  $k \leq d \leq n$  and, for some  $\beta \in (0, 1]$ ,  $\|Dg\|_{C^{k+1+\beta}} < \infty$ . In addition, suppose that for some  $\delta > 0$   $\sigma(\Sigma) \subset [\delta, \frac{1}{\delta}]$ . Then, for some constant  $C > 0$ ,

$$\begin{aligned} & \|\mathbb{E}_{\Sigma} \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| \\ & \leq C^{k^2} \frac{\log^2(2/\delta)}{\delta} \|Dg\|_{C^{k+1+\beta}} (\|\Sigma\| \vee 1)^{k+3/2} \|\Sigma\| \left(\frac{d}{n}\right)^{(k+1+\beta)/2}. \end{aligned}$$

# Sketch of the proof: bounding partial derivatives

- To compute  $\frac{\partial^k \varphi}{\partial t_1 \dots \partial t_k}$ , we derive formulas for partial derivatives of operator-valued function  $h(S(t_1, \dots, t_k))$ ,  $h = Dg$ . Recall that

$$R := R(t_1, \dots, t_k) = V_1(t_1) \dots V_k(t_k),$$

$$L := L(t_1, \dots, t_k) = V_k(t_k) \dots V_1(t_1) = R^*,$$

$$S := S(t_1, \dots, t_k) = L(t_1, \dots, t_k) \Sigma R(t_1, \dots, t_k).$$

- Given  $T = \{t_{i_1}, \dots, t_{i_m}\} \subset \{t_1, \dots, t_k\}$ , let  $\partial_T S := \frac{\partial^m S(t_1, \dots, t_k)}{\partial t_{i_1} \dots \partial t_{i_m}}$  (similar notations are used for partial derivatives of  $h(S)$ , etc.).
- Let  $\mathcal{D}_{j,T}$  be the set of all partitions  $(\Delta_1, \dots, \Delta_j)$  of  $T \subset \{t_1, \dots, t_k\}$  with non-empty sets  $\Delta_i$ ,  $i = 1, \dots, j$  (partitions with different order of  $\Delta_1, \dots, \Delta_j$  being identical).
- For  $\Delta = (\Delta_1, \dots, \Delta_j) \in \mathcal{D}_{j,T}$ , set  $\partial_\Delta S = (\partial_{\Delta_1} S, \dots, \partial_{\Delta_j} S)$ .
- Denote  $\mathcal{D}_T := \bigcup_{j=1}^{|T|} \mathcal{D}_{j,T}$ .
- For  $\Delta = (\Delta_1, \dots, \Delta_j) \in \mathcal{D}_T$ , set  $j_\Delta := j$ .

## Lemma

Suppose, for some  $m \leq k$ ,  $h = Dg \in L_\infty(\mathcal{C}_+^d; \mathcal{B}_{sa}(\mathbb{R}^d))$  is  $m$  times continuously differentiable with derivatives  $D^j h, j \leq m$ . Then the function  $[0, 1]^k \ni (t_1, \dots, t_k) \mapsto h(S(t_1, \dots, t_k))$  is  $m$  times continuously differentiable and for any  $T \subset \{t_1, \dots, t_k\}$  with  $|T| = m$

$$\partial_T h(S) = \sum_{\Delta \in \mathcal{D}_T} D^{j_\Delta} h(S)(\partial_\Delta S) = \sum_{j=1}^m \sum_{\Delta \in \mathcal{D}_{j,T}} D^j h(S)(\partial_\Delta S).$$



# Sketch of the proof: bounding partial derivatives

Denote

$$\delta_i := \|W_i - I\|, i = 1, \dots, k.$$

## Lemma

For all  $T \subset \{t_1, \dots, t_k\}$ ,

$$\|\partial_T R\| \leq \prod_{t_i \in T} \frac{\delta_i}{1 + \delta_i} \prod_{i=1}^k (1 + \delta_i),$$

$$\|\partial_T L\| \leq \prod_{t_i \in T} \frac{\delta_i}{1 + \delta_i} \prod_{i=1}^k (1 + \delta_i),$$

$$\|\partial_T S\| \leq 2^k \|\Sigma\| \prod_{t_i \in T} \frac{\delta_i}{1 + \delta_i} \prod_{i=1}^k (1 + \delta_i)^2.$$

## Lemma

Suppose that, for some  $0 \leq m \leq k$ ,  $h = Dg \in L_\infty(\mathcal{C}_+^d; \mathcal{B}_{sa}(\mathbb{R}^d))$  is  $m$  times differentiable with uniformly bounded continuous derivatives  $D^j h, j = 1, \dots, m$ . Then for all  $T \subset \{t_1, \dots, t_k\}$  with  $|T| = m$

$$\|\partial_T h(\mathbf{S})\| \leq 2^{m(k+m+1)} \max_{0 \leq j \leq m} \|D^j h\|_{L_\infty} (\|\Sigma\|^m \vee 1) \prod_{i=1}^k (1 + \delta_i)^{2m} \prod_{t_j \in T} \frac{\delta_j}{1 + \delta_j}$$

## Lemma

$$\left\| \frac{\partial^k \varphi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} \right\|$$
$$\leq 3^k 2^{k(2k+1)} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \prod_{i=1}^k \delta_i (1 + \delta_i)^{2k+1},$$

where  $\delta_i := \|W_i - I\|$ .

# Sketch of the proof: bound on $\|\mathcal{B}^k \mathcal{D}g(\Sigma)\|$

$$\begin{aligned} \|\mathcal{B}^k \mathcal{D}g(\Sigma)\| &= \|\mathcal{D}\mathcal{B}^k g(\Sigma)\| \\ &\leq \mathbb{E} \int_0^1 \dots \int_0^1 \left\| \frac{\partial^k \varphi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} \right\| dt_1 \dots dt_k \\ &\leq 3^k 2^{k(2k+1)} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \mathbb{E} \prod_{i=1}^k \delta_i (1 + \delta_i)^{2k+1} \\ &\leq 3^k 2^{k(2k+1)} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \\ &\quad \times \left( \mathbb{E} \|W - I\| (1 + \|W - I\|)^{2k+1} \right)^k \\ &\leq C^{k^2} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \left( \frac{d}{n} \right)^{k/2}. \end{aligned}$$

## Part 4.

# Asymptotic Efficiency

Let  $X, X_1, \dots, X_n$  be i.i.d. Gaussian vectors with values in  $\mathbb{R}^d$ , with  $\mathbb{E}X = 0$  and with covariance operator  $\Sigma = \mathbb{E}(X \otimes X) \in \mathcal{C}_+^d$ .

- Given a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$  and a linear operator  $B : \mathbb{R}^d \mapsto \mathbb{R}^d$  with  $\|B\|_1 \leq 1$ , estimate  $\langle f(\Sigma), B \rangle$  based on  $X_1, \dots, X_n$ .
- More precisely, we are interested in finding **asymptotically efficient** estimators of  $\langle f(\Sigma), B \rangle$  with  $\sqrt{n}$ -convergence rate in the case when  $d = d_n \rightarrow \infty$ .
- Suppose  $d_n \leq n^\alpha$  for some  $\alpha > 0$ . Is there  $s(\alpha)$  such that for all  $s > s(\alpha)$  and for all functions  $f$  of smoothness  $s$ , asymptotically efficient estimation is possible?

# Sample Covariance Operator

- Let

$$\hat{\Sigma} := n^{-1} \sum_{j=1}^n X_j \otimes X_j$$

be the sample covariance based on  $(X_1, \dots, X_n)$ .

- Let

$$\mathcal{S}_{a,d} := \left\{ \Sigma \in \mathcal{C}_+^d : a^{-1} I_d \preceq \Sigma \preceq a I_d \right\}, a > 1.$$

- If  $\Sigma \in \mathcal{S}_{a,d}$ , then

$$\mathbb{E} \|\hat{\Sigma} - \Sigma\| \asymp_a \|\Sigma\| \left( \sqrt{\frac{d}{n}} V \frac{d}{n} \right)$$

and, for all  $t \geq 1$  with probability at least  $1 - e^{-t}$ ,

$$\|\hat{\Sigma} - \Sigma\| \lesssim_a \|\Sigma\| \left( \sqrt{\frac{d}{n}} V \frac{d}{n} V \sqrt{\frac{t}{n}} V \frac{t}{n} \right).$$

- Let  $f \in C^1(\mathbb{R})$  and let  $f^{[1]}(\lambda, \mu)$  be the Loewner kernel:

$$f^{[1]}(\lambda, \mu) := \frac{f(\lambda) - f(\mu)}{\lambda - \mu}, \quad \lambda \neq \mu; \quad f^{[1]}(\lambda, \lambda) := f'(\lambda).$$

- $A \mapsto f(A)$  is Fréchet differentiable at  $A = \sum_{\lambda \in \sigma(A)} \lambda P_\lambda$  with derivative

$$Df(A; H) = \sum_{\lambda, \mu \in \sigma(A)} f^{[1]}(\lambda, \mu) P_\lambda H P_\mu.$$



# Assumptions and Notations

- Let

$$\sigma_f^2(\Sigma; B) := 2\|\Sigma^{1/2}Df(\Sigma; B)\Sigma^{1/2}\|_2^2.$$

- **Loss functions.** Let  $\mathcal{L}$  be the class of functions  $\ell : \mathbb{R} \mapsto \mathbb{R}_+$  such that
  - $\ell(0) = 0$
  - $\ell(-t) = \ell(t), t \in \mathbb{R}$
  - $\ell$  is convex and nondecreasing on  $\mathbb{R}_+$
  - For some  $c > 0$ ,  $\ell(t) = O(e^{ct})$  as  $t \rightarrow \infty$
- Suppose that
  - **A.1.**  $d_n \geq 3 \log n$
  - **A.2.** for some  $\alpha \in (0, 1)$ ,  $d_n \leq n^\alpha$
  - **A.3.** For some  $s > \frac{1}{1-\alpha}$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ .

# Efficient Estimation of $\langle f(\Sigma), B \rangle$

## Theorem

Under assumptions **A.1-A.3**, there exists an estimator  $h(\hat{\Sigma})$  such that for all  $\sigma_0 > 0$

$$\sup_{\Sigma \in \mathcal{S}_{a,d_n}, \sigma_f(\Sigma; B) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_{\Sigma} \left\{ \frac{n^{1/2} \left( \langle h(\hat{\Sigma}), B \rangle - \langle f(\Sigma), B \rangle \right)}{\sigma_f(\Sigma, B)} \leq x \right\} - \Phi(x) \right| \rightarrow 0$$

and, for all  $\ell \in \mathcal{L}$ ,

$$\sup_{\Sigma \in \mathcal{S}_{a,d_n}, \sigma_f(\Sigma; B) \geq \sigma_0} \left| \mathbb{E}_{\Sigma} \ell \left( \frac{n^{1/2} \left( \langle h(\hat{\Sigma}), B \rangle - \langle f(\Sigma), B \rangle \right)}{\sigma_f(\Sigma, B)} \right) - \mathbb{E} \ell(Z) \right| \rightarrow 0.$$

# Efficient Estimation of $\langle f(\Sigma), B \rangle$ : A Lower Bound

## Theorem

Let, for some  $s \in (1, 2]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ . Suppose  $d_n \geq 1$ ,  $a > 1$ ,  $\sigma_0 > 0$  are such that, for some  $1 < a' < a$  and  $\sigma'_0 > \sigma_0$  and for all large enough  $n$ ,

$$\left\{ \Sigma \in \mathcal{S}_{a,d_n}, \sigma_f(\Sigma; B) \geq \sigma_0 \right\} \neq \emptyset.$$

Then, the following bound holds:

$$\liminf_n \inf_{T_n} \sup_{\Sigma \in \mathcal{S}_{a,d_n}, \sigma_f(\Sigma; B) \geq \sigma_0} \frac{n \mathbb{E}_{\Sigma} \left( T_n(X_1, \dots, X_n) - \langle f(\Sigma), B \rangle \right)^2}{\sigma_f^2(\Sigma; B)} \geq 1,$$

where the infimum is taken over all sequences of estimators  $T_n = T_n(X_1, \dots, X_n)$ .

# Lower bound: sketch of the proof

- Let

$$\Sigma_t := \Sigma_0 + \frac{tH}{\sqrt{n}} \text{ and } \mathcal{S}_{c,n}(\Sigma_0, H) := \{\Sigma_t : t \in [-c, c]\},$$

where

$$H := \Sigma_0 Df(\Sigma_0; B) \Sigma_0.$$

- For all large enough  $n$ ,

$$\mathcal{S}_{c,n}(\Sigma_0, H) \subset \mathcal{S}_{a,d_n} \cap \{\Sigma : \sigma_f(\Sigma, B) \geq \sigma_0\}.$$

# Lower bound: sketch of the proof

- Consider the following parametric model:

$$X_1, \dots, X_n \text{ i.i.d. } \sim N(0; \Sigma_t), t \in [-c, c].$$

- The goal is to estimate the function

$$\varphi(t) := \langle f(\Sigma_t), B \rangle.$$

which is continuously differentiable with derivative

$$\varphi'(t) = \frac{1}{\sqrt{n}} \langle Df(\Sigma_t; H), B \rangle, t \in [-c, c].$$

- The Fisher information:**

$$\begin{aligned} I_n(t) &= nI(t) = \frac{1}{2} \langle I(\Sigma_t)H, H \rangle = \frac{1}{2} \langle (\Sigma_t^{-1} \otimes \Sigma_t^{-1})H, H \rangle \\ &= \frac{1}{2} \langle \Sigma_t^{-1} H \Sigma_t^{-1}, H \rangle = \frac{1}{2} \text{tr}(\Sigma_t^{-1} H \Sigma_t^{-1} H) = \frac{1}{2} \|\Sigma_t^{-1/2} H \Sigma_t^{-1/2}\|_2^2. \end{aligned}$$

# Lower bound: sketch of the proof

- Let  $\pi_c(t) := \frac{1}{c}\pi\left(\frac{t}{c}\right)$  for a smooth density  $\pi$  on  $[-1, 1]$  with  $\pi(1) = \pi(-1) = 0$  and  $J_\pi := \int_{-1}^1 \frac{(\pi'(t))^2}{\pi(t)} dt < \infty$ .
- **Van Trees Inequality:** for any estimator  $T(X_1, \dots, X_n)$ ,

$$\begin{aligned} & \sup_{t \in [-c, c]} \mathbb{E}_t(T_n(X_1, \dots, X_n) - \varphi(t))^2 \\ & \geq \int_{-c}^c \mathbb{E}_t(T_n(X_1, \dots, X_n) - \varphi(t))^2 \pi_c(t) dt \geq \frac{\left(\int_{-c}^c \varphi'(t) \pi_c(t) dt\right)^2}{\int_{-c}^c I_n(t) \pi_c(t) dt + J_{\pi_c}}. \end{aligned}$$

# Lower bound: sketch of the proof

Denote  $\sigma^2(t) := \sigma_f^2(\Sigma_t; B)$ ,  $t \in [-c, c]$ . Then

$$\begin{aligned} & \sup_{\Sigma \in \mathcal{S}_{a,d_n}, \sigma_f(\Sigma; B) \geq \sigma_0} \frac{n \mathbb{E}_{\Sigma} (T_n(X_1, \dots, X_n) - \langle f(\Sigma), B \rangle)^2}{\sigma_f^2(\Sigma; B)} \\ & \geq \sup_{t \in [-c, c]} \frac{n \mathbb{E}_t (T_n(X_1, \dots, X_n) - \varphi(t))^2}{\sigma^2(t)} \geq \frac{1 - \gamma_{n,c}(f; B; a; \sigma_0)}{1 + \frac{\lambda_{n,c}(f; B; a)}{\sigma_0^2}}, \end{aligned}$$

where

$$\gamma_{n,c}(f; B; a; \sigma_0) := \frac{\frac{3a^3 \|f'\|_{L_{\infty}}^3 \|B\|_1^3 c}{\sqrt{n}} + \frac{4a^{2s} \|f\|_{B_{\infty,1}^s} \|f'\|_{L_{\infty}}^s \|B\|_1^{1+s} c^{s-1}}{n^{(s-1)/2}} + \frac{J_{\pi}}{c^2}}{\frac{1}{4}\sigma_0^2 + \frac{3a^3 \|f'\|_{L_{\infty}}^3 \|B\|_1^3 c}{\sqrt{n}} + \frac{J_{\pi}}{c^2}}$$

and

$$\lambda_{n,c}(f; B; a) := \frac{6ca^3 \|f'\|_{L_{\infty}}^3 \|B\|_1^3}{n^{1/2}} + \frac{24c^{s-1} a^{2s} \|f'\|_{L_{\infty}}^s \|f\|_{B_{\infty,1}^s} \|B\|_1^{s+1}}{n^{(s-1)/2}}.$$

# Construction of an Asymptotically Efficient Estimator

- If  $d_n \leq n^\alpha$  for some  $\alpha \in (0, 1/2)$  and  $s > \frac{1}{1-\alpha}$ , then the plug-in estimator  $\langle f(\hat{\Sigma}), B \rangle$  is asymptotically efficient with convergence rate  $\sqrt{n}$ .
- If  $d_n \geq n^\alpha$  for some  $\alpha \geq 1/2$ , then the plug-in estimator  $\langle f(\hat{\Sigma}), B \rangle$ , for “generic” smooth functions  $f$  has a large bias (larger than  $n^{-1/2}$ ) and it is not even  $\sqrt{n}$  consistent.
- In the last case, the crucial problem is to construct an estimator with a reduced bias.



# Bounds on the Remainder of Differentiation

## Lemma

Let  $S_f(A; H) = f(A + H) - f(A) - (Df)(A; H)$  be the remainder of differentiation. If, for some  $s \in [1, 2]$ ,  $f \in B_{\infty,1}^s(\mathbb{R})$ , then the following bounds hold:

$$\|S_f(A; H)\| \leq 2^{3-s} \|f\|_{B_{\infty,1}^s} \|H\|^s$$

and

$$\|S_f(A; H) - S_f(A; H')\| \leq 2^{1+s} \|f\|_{B_{\infty,1}^s} (\|H\| \vee \|H'\|)^{s-1} \|H' - H\|.$$

The proof is based on Littlewood-Paley decomposition of  $f$  and on operator versions of Bernstein inequalities for entire functions of exponential type (as in the work by [Peller \(1985, 2006\)](#), [Aleksandrov and Peller \(2016\)](#) on operator Lipschitz and operator differentiable functions).

# Perturbation Theory: Application to Functions of Sample Covariance (The Delta Method)

- $$\langle f(\hat{\Sigma}) - f(\Sigma), B \rangle = \langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle + \langle S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$$
- The linear term  $\langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$  is of the order  $O(n^{-1/2})$  and  $n^{1/2}\langle Df(\Sigma; \hat{\Sigma} - \Sigma), B \rangle$  is close in distribution to  $N(0; \sigma_f^2(\Sigma; B))$ .
- For  $s \in (1, 2]$ ,  $\|S_f(\Sigma; \hat{\Sigma} - \Sigma)\| \lesssim \|f\|_{B_{\infty,1}^s} \|\hat{\Sigma} - \Sigma\|^s$ , implying that

$$\begin{aligned} |\langle S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle| &\leq \|B\|_1 \|S_f(\Sigma; \hat{\Sigma} - \Sigma)\| \\ &= O\left(\left(\frac{d}{n}\right)^{s/2}\right) = O(n^{(1-\alpha)s/2}) = o(n^{-1/2}) \end{aligned}$$

and, similarly,

$$|\langle \mathbb{E}f(\hat{\Sigma}) - f(\Sigma), B \rangle| = |\langle \mathbb{E}S_f(\Sigma; \hat{\Sigma} - \Sigma), B \rangle| = o(n^{-1/2})$$

provided that  $s > \frac{1}{1-\alpha}$ ,  $\alpha \in (0, 1/2)$ . In this case,  $h(\hat{\Sigma}) = f(\hat{\Sigma})$ .

- Our next goal is to find an estimator  $g(\hat{\Sigma})$  of  $f(\Sigma)$  with a small bias  $\mathbb{E}_{\Sigma}g(\hat{\Sigma}) - f(\Sigma)$  (of the order  $o(n^{-1/2})$ ) and such that

$$n^{1/2}(\langle g(\hat{\Sigma}), B \rangle - \langle \mathbb{E}_{\Sigma}g(\hat{\Sigma}), B \rangle)$$

is asymptotically normal.

- To this end, one has to find a sufficiently smooth approximate solution of the equation

$$\mathbb{E}_{\Sigma}g(\hat{\Sigma}) = f(\Sigma), \Sigma \in \mathcal{C}_+^d.$$

# Wishart Operators and Bias Reduction



$$\mathcal{T}g(\Sigma) := \mathbb{E}_{\Sigma}g(\hat{\Sigma}) = \int_{\mathcal{C}_+^d} g(V)P(\Sigma; dV), \Sigma \in \mathcal{C}_+^d$$

- To find an estimator of  $f(\Sigma)$  with a small bias, one needs to solve (approximately) the following integral equation (“the Wishart equation”)

$$\mathcal{T}g(\Sigma) = f(\Sigma), \Sigma \in \mathcal{C}_+^d.$$

- Bias operator:  $\mathcal{B} := \mathcal{T} - \mathcal{I}$ .
- Formally, the solution of the Wishart equation is given by Neumann series:

$$g(\Sigma) = (\mathcal{I} + \mathcal{B})^{-1}f(\Sigma) = (\mathcal{I} - \mathcal{B} + \mathcal{B}^2 - \dots)f(\Sigma)$$

# Wishart Operators and Bias Reduction

- Given a smooth function  $f : \mathbb{R} \mapsto \mathbb{R}$ , define

$$f_k(\Sigma) := \sum_{j=0}^k (-1)^j \mathcal{B}^j f(\Sigma) := f(\Sigma) + \sum_{j=1}^k (-1)^j \mathcal{B}^j f(\Sigma)$$

- Then

$$\mathbb{E}_{\Sigma} f_k(\hat{\Sigma}) - f(\Sigma) = (\mathcal{I} + \mathcal{B})f_k(\Sigma) - f(\Sigma) = (-1)^k \mathcal{B}^{k+1} f(\Sigma).$$

- Asymptotically efficient estimator is  $h(\hat{\Sigma}) = f_k(\hat{\Sigma})$ , where  $k$  is an integer such that, for some  $\beta \in (0, 1]$ ,  $\frac{1}{1-\alpha} < k + 1 + \beta \leq s$ .



$$\mathcal{T}g(\Sigma) := \mathbb{E}_{\Sigma}g(\hat{\Sigma}) = \int_{\mathcal{C}_+^d} g(V)P(\Sigma; dV), \Sigma \in \mathcal{C}_+^d$$



$$\mathcal{T}^k g(\Sigma) = \mathbb{E}_{\Sigma}g(\hat{\Sigma}^{(k)}), \Sigma \in \mathcal{C}_+^d,$$

where

$$\hat{\Sigma}^{(0)} = \Sigma \rightarrow \hat{\Sigma}^{(1)} = \hat{\Sigma} \rightarrow \hat{\Sigma}^{(2)} \rightarrow \dots$$

is a Markov chain in  $\mathcal{C}_+^d$  with transition probability kernel  $P$ .

- Note that  $\hat{\Sigma}^{(j+1)}$  is the sample covariance based on  $n$  i.i.d. observations  $\sim N(0; \hat{\Sigma}^{(j)})$  (conditionally on  $\hat{\Sigma}^{(j)}$ )
- Conditionally on  $\hat{\Sigma}^{(j)}$ , with a “high probability”,

$$\|\hat{\Sigma}^{(j+1)} - \hat{\Sigma}^{(j)}\| \lesssim \|\hat{\Sigma}^{(j)}\| \sqrt{\frac{d}{n}}$$

- $k$ -th order difference along the Markov chain:

$$\mathcal{B}^k f(\Sigma) = (\mathcal{T} - \mathcal{I})^k f(\Sigma) = \sum_{j=0}^k (-1)^{k-j} \binom{k}{j} \mathcal{T}^j f(\Sigma)$$

$$= \mathbb{E}_{\Sigma} \sum_{j=0}^k (-1)^{k-j} \binom{k}{j} f(\hat{\Sigma}^{(j)})$$

- $\sum_{j=0}^k (-1)^{k-j} \binom{k}{j} f(\hat{\Sigma}^{(j)})$  is the  $k$ -th order difference of  $f$  along the trajectory of the Bootstrap Chain. For  $f$  of smoothness  $k$ , it should be of the order  $\left(\sqrt{\frac{d}{n}}\right)^k$ .

## Theorem

Suppose that  $f \in B_{\infty,1}^k(\mathbb{R})$  and that  $k \leq d \leq n$ . Then, for some  $C > 1$ ,

$$\|\mathcal{B}^k f(\Sigma)\| \leq C^{k^2} \|f\|_{B_{\infty,1}^k} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{k/2}.$$



# Bounds on the bias of $f_k(\hat{\Sigma})$

## Corollary

Suppose  $f \in B_{\infty,1}^{k+1}(\mathbb{R})$  and  $k+1 \leq d \leq n$ . Then, for some  $C > 1$ ,

$$\|\mathbb{E}_{\Sigma} f_k(\hat{\Sigma}) - f(\Sigma)\| \leq C^{(k+1)^2} \|f\|_{B_{\infty,1}^{k+1}} (\|\Sigma\|^{k+2} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{(k+1)/2}.$$

If, for some  $\alpha \in (1/2, 1)$ ,  $2 \log n \leq d \leq n^\alpha$  and  $k > \frac{\alpha}{1-\alpha}$ , then

$$\|\mathbb{E}_{\Sigma} f_k(\hat{\Sigma}) - f(\Sigma)\| = o(n^{-1/2}).$$

# “Lifting” Operator $\mathcal{D}$ and Reduction to Orthogonally Invariant Functions

- Define the following differential operator (“lifting” operator):

$$\mathcal{D}g(\Sigma) := \Sigma^{1/2} Dg(\Sigma) \Sigma^{1/2}$$

acting on continuously differentiable functions in  $\mathcal{C}_+^d$ .

- Suppose  $f(x) = x\psi'(x)$
- Let  $g(\Sigma) := \text{tr}(\psi(\Sigma))$
- $g$  is orthogonally invariant function on  $\mathcal{C}_+^d$
- $Dg(\Sigma) = \psi'(\Sigma)$
- $\mathcal{D}g(\Sigma) = f(\Sigma)$

# An Integral Representation for Operator $\mathcal{B}^k$

- Linear interpolation between  $I$  and  $W_1^{1/2}, \dots, W_k^{1/2}$  (i.i.d. copies of  $W = n^{-1} \sum_{j=1}^n Z_j \otimes Z_j$ ):

$$V_j(t_j) := I + t_j(W_j^{1/2} - I), t_j \in [0, 1], 1 \leq j \leq k.$$

- For all  $j = 1, \dots, k, t_j \in [0, 1], V_j(t_j) \in \mathcal{C}_+^d$ .



$$R := R(t_1, \dots, t_k) = V_1(t_1) \dots V_k(t_k),$$

$$L := L(t_1, \dots, t_k) = V_k(t_k) \dots V_1(t_1) = R^*,$$

$$S := S(t_1, \dots, t_k) = L(t_1, \dots, t_k) \Sigma R(t_1, \dots, t_k), (t_1, \dots, t_k) \in [0, 1]^k.$$

- Let

$$\varphi(t_1, \dots, t_k) := \Sigma^{1/2} R(t_1, \dots, t_k) Dg(S(t_1, \dots, t_k)) L(t_1, \dots, t_k) \Sigma^{1/2}.$$

## Proposition

Suppose  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 1$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 1$ . Then the function  $\varphi$  is  $k$  times continuously differentiable in  $[0, 1]^k$  and

$$\mathcal{B}^k \mathcal{D}g(\Sigma) = \mathbb{E} \int_0^1 \dots \int_0^1 \frac{\partial^k \varphi(t_1, \dots, t_k)}{\partial t_1 \dots \partial t_k} dt_1 \dots dt_k, \Sigma \in \mathcal{C}_+^d.$$

# A bound on $\mathcal{B}^k \mathcal{D}g(\Sigma)$

## Theorem

Suppose that  $k \leq d \leq n$  and that  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 1$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 1$ . Then, for some  $C > 1$ ,

$$\|\mathcal{B}^k \mathcal{D}g(\Sigma)\| \leq C^{k^2} \max_{1 \leq j \leq k+1} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+1} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{k/2}.$$

# Bounds on the bias of $\mathcal{D}g_k(\hat{\Sigma})$

## Corollary

Suppose that  $k + 1 \leq d \leq n$  and that  $g \in L_\infty^O(\mathcal{C}_+^d)$  is  $k + 2$  times continuously differentiable function with uniformly bounded derivatives  $D^j g, j = 1, \dots, k + 2$ . Then, for some  $C > 1$ ,

$$\begin{aligned} & \|\mathbb{E}_\Sigma \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| \\ & \leq C^{(k+1)^2} \max_{1 \leq j \leq k+2} \|D^j g\|_{L_\infty} (\|\Sigma\|^{k+2} \vee \|\Sigma\|) \left(\frac{d}{n}\right)^{(k+1)/2}. \end{aligned}$$

If, for some  $\alpha \in (1/2, 1)$ ,  $2 \log n \leq d \leq n^\alpha$  and  $k > \frac{\alpha}{1-\alpha}$ , then

$$\|\mathbb{E}_\Sigma \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| = o(n^{-1/2}).$$

# Further bounds on the bias of $\mathcal{D}g_k(\hat{\Sigma})$

## Theorem

Suppose  $g \in L_{\infty}^O(\mathcal{C}_+^d)$  is  $k + 2$  times continuously differentiable for some  $k \leq d \leq n$  and, for some  $\beta \in (0, 1]$ ,  $\|Dg\|_{C^{k+1+\beta}} < \infty$ . In addition, suppose that for some  $\delta > 0$   $\sigma(\Sigma) \subset [\delta, \frac{1}{\delta}]$ . Then, for some constant  $C > 0$ ,

$$\begin{aligned} & \|\mathbb{E}_{\Sigma} \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma)\| \\ & \leq C k^2 \frac{\log^2(2/\delta)}{\delta} \|Dg\|_{C^{k+1+\beta}} (\|\Sigma\| \vee 1)^{k+3/2} \|\Sigma\| \left(\frac{d}{n}\right)^{(k+1+\beta)/2}. \end{aligned}$$

# Further details of the proof

- We use the representation

$$\begin{aligned} \langle \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle &= \langle D\mathcal{D}g_k(\Sigma)(\hat{\Sigma} - \Sigma), B \rangle \\ &+ \mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) + \langle \mathbb{E}\mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle, \end{aligned}$$

where  $\mathfrak{d}_k(\Sigma) := \langle \mathcal{D}g_k(\Sigma), B \rangle$ .

- We control the bias  $\langle \mathbb{E}\mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle$  as follows:

$$|\langle \mathbb{E}\mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle| = O\left(\frac{d}{n}\right)^{(k+1+\beta)/2} = o(n^{-1/2})$$

provided that  $d \leq n^\alpha$ ,  $\alpha \in (0, 1)$  and  $s \geq k + 1 + \beta > \frac{1}{1-\alpha}$ .

- We also need to prove concentration of

$$\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma).$$



# Lipschitz Condition for the Remainder of Taylor Expansion of $Dg_k(\Sigma)$

## Lemma

Suppose that, for some  $k \leq d$ ,  $g$  is  $k + 2$  times continuously differentiable and, for some  $\beta \in (0, 1]$ ,  $\|Dg\|_{C^{k+1+\beta}} < \infty$ . Suppose also that  $g \in L_\infty^O(C_+^d)$ ,  $d \leq n/2$  and that, for some  $\delta \in (0, 1)$ ,  $\sigma(\Sigma) \subset [\delta, \frac{1}{\delta}]$ . Denote

$$\gamma_{\beta,k}(\Sigma; u) := (\|\Sigma\| \vee u \vee 1)^{k+1/2} (u \vee u^\beta), \quad u > 0, \beta \in [0, 1], k \geq 1.$$

Then, for some constant  $C \geq 1$  and for all  $H, H' \in \mathcal{B}_{\text{sa}}(\mathbb{R}^d)$

$$\begin{aligned} & \|S_{Dg_k}(\Sigma; H') - S_{Dg_k}(\Sigma; H)\| \\ & \leq C^{k^2} \frac{\log^2(2/\delta)}{\delta} \|Dg\|_{C^{k+1+\beta}} \gamma_{\beta,k}(\Sigma; \|H\| \vee \|H'\|) \|H' - H\|. \end{aligned}$$

# Sketch of the proof of the Lemma

- Let  $\gamma : \mathbb{R} \mapsto \mathbb{R}$  be a  $C^\infty$  function such that:

$$0 \leq \gamma(u) \leq \sqrt{u}, u \geq 0, \gamma(u) = \sqrt{u}, u \in \left[\delta, \frac{1}{\delta}\right],$$
$$\text{supp}(\gamma) \subset \left[\frac{\delta}{2}, \frac{2}{\delta}\right] \text{ and } \|\gamma\|_{B_{\infty,1}^1} \lesssim \frac{\log(2/\delta)}{\sqrt{\delta}}.$$

- For instance, one can take  $\gamma(u) := \lambda(u/\delta)\sqrt{u}(1 - \lambda(\delta u/2))$ , where  $\lambda$  is a  $C^\infty$  non-decreasing function with values in  $[0, 1]$ ,  $\lambda(u) = 0, u \leq 1/2$  and  $\lambda(u) = 1, u \geq 1$ .

# Further Integral Representations



$$\begin{aligned} \phi(t_1, \dots, t_k; \mathbf{s}_1, \mathbf{s}_2) := \\ \gamma(\bar{\Sigma}(\mathbf{s}_1, \mathbf{s}_2)) R(t_1, \dots, t_k) Dg \left( L(t_1, \dots, t_k) \bar{\Sigma}(\mathbf{s}_1, \mathbf{s}_2) R(t_1, \dots, t_k) \right) \\ L(t_1, \dots, t_k) \gamma(\bar{\Sigma}(\mathbf{s}_1, \mathbf{s}_2)), \end{aligned}$$

where

$$\bar{\Sigma}(\mathbf{s}_1, \mathbf{s}_2) = \Sigma + \mathbf{s}_1 H + \mathbf{s}_2 (H' - H), \mathbf{s}_1, \mathbf{s}_2 \in \mathbb{R}.$$

- Note that

$$\varphi(t_1, \dots, t_k) = \phi(t_1, \dots, t_k, \mathbf{0}, \mathbf{0}).$$

# Further Integral Representations

- $$B_k(\Sigma) := \mathcal{B}^k \mathcal{D}g(\Sigma), \quad D_k(\Sigma) := \mathcal{D}g_k(\Sigma).$$

- $$D_k(\Sigma) = \sum_{j=0}^k (-1)^j B_j(\Sigma)$$

- $$B_k(\Sigma) := \mathbb{E} \int_0^1 \dots \int_0^1 \frac{\partial^k \phi(t_1, \dots, t_k, 0, 0)}{\partial t_1 \dots \partial t_k} dt_1 \dots dt_k$$

# Further Integral Representations

- $$\begin{aligned} S_{B_k}(\Sigma; H') - S_{B_k}(\Sigma; H) \\ = DB_k(\Sigma + H; H' - H) - DB_k(\Sigma; H' - H) + S_{B_k}(\Sigma + H; H' - H) \end{aligned}$$

- $$\begin{aligned} DB_k(\Sigma + H; H' - H) - DB_k(\Sigma; H' - H) \\ = \mathbb{E} \int_0^1 \cdots \int_0^1 \left[ \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 1, 0)}{\partial t_1 \dots \partial t_k \partial s_2} \right. \\ \left. - \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 0, 0)}{\partial t_1 \dots \partial t_k \partial s_2} \right] dt_1 \dots dt_k \end{aligned}$$



$$\begin{aligned} & S_{B_k}(\Sigma + H; H' - H) \\ &= \mathbb{E} \int_0^1 \dots \int_0^1 \int_0^1 \left[ \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 1, \mathbf{s}_2)}{\partial t_1 \dots \partial t_k \partial \mathbf{s}_2} \right. \\ & \quad \left. - \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 1, 0)}{\partial t_1 \dots \partial t_k \partial \mathbf{s}_2} \right] d\mathbf{s}_2 dt_1 \dots dt_k. \end{aligned}$$

- The rest of the proof is based on bounding the partial derivatives involved in the integral representations, such as, for instance,

$$\begin{aligned} & \left\| \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 1, 0)}{\partial t_1 \dots \partial t_k \partial s_2} - \frac{\partial^{k+1} \phi(t_1, \dots, t_k, 0, 0)}{\partial t_1 \dots \partial t_k \partial s_2} \right\| \\ & \leq C^{k^2} \frac{\log^2(2/\delta)}{\delta} \|Dg\|_{C^{k+1+\beta}} (\|\Sigma\| + \|H\|)^{k+1/2} \vee 1) \\ & \prod_{i=1}^k \delta_i (1 + \delta_i)^{2k+5} (\|H\| \vee \|H\|^\beta) \|H' - H\|, \end{aligned}$$

where  $\delta_i := \|W_i - I\|$ .

# Concentration of the Remainder: A More General Version

## Assumption

Assume that, for all  $\Sigma \in \mathcal{C}_+(\mathbb{H})$ ,  $H, H' \in \mathcal{B}_{sa}(\mathbb{H})$ ,

$$|S_h(\Sigma; H') - S_h(\Sigma; H)| \leq \eta(\Sigma; \|H\| \vee \|H'\|) \|H' - H\|,$$

where  $0 < \delta \mapsto \eta(\Sigma; \delta)$  is a nondecreasing function of the following form:

$$\eta(\Sigma; \delta) := \eta_1(\Sigma) \delta^{\alpha_1} \bigvee \cdots \bigvee \eta_m(\Sigma) \delta^{\alpha_m},$$

for given nonnegative functions  $\eta_1, \dots, \eta_m$  on  $\mathcal{C}_+(\mathbb{H})$  and given positive numbers  $\alpha_1, \dots, \alpha_m$ .



# Concentration of the Remainder: A More General Version

## Theorem

For all  $t \geq 1$  with probability at least  $1 - e^{-t}$ ,

$$\begin{aligned} & |S_h(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_h(\Sigma; \hat{\Sigma} - \Sigma)| \\ & \lesssim_{\eta} \eta(\Sigma; \delta_n(\Sigma; t)) \left( \sqrt{\|\Sigma\|} + \sqrt{\delta_n(\Sigma; t)} \right) \sqrt{\|\Sigma\|} \sqrt{\frac{t}{n}}, \end{aligned}$$

where

$$\delta_n(\Sigma; t) := \|\Sigma\| \left( \sqrt{\frac{\mathbf{r}(\Sigma)}{n}} \vee \frac{\mathbf{r}(\Sigma)}{n} \vee \sqrt{\frac{t}{n}} \vee \frac{t}{n} \right).$$

# Concentration Inequality for the Remainder of Taylor Expansion of $\langle Dg_k(\Sigma), B \rangle$

## Lemma

With probability at least  $1 - e^{-t}$ ,

$$\begin{aligned} & |S_{\partial_k}(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}S_{\partial_k}(\Sigma; \hat{\Sigma} - \Sigma)| \\ & \leq C^{k^2} \tilde{\Lambda}_{k,\beta}(g; \Sigma; B) \gamma_{\beta,k}(\Sigma; \bar{\delta}_n(\Sigma; t)) \left( \sqrt{\|\Sigma\|} + \sqrt{\bar{\delta}_n(\Sigma; t)} \right) \sqrt{\|\Sigma\|} \sqrt{\frac{t}{n}}, \end{aligned}$$

where

$$\bar{\delta}_n(\Sigma; t) := \|\Sigma\| \left( \sqrt{\frac{d}{n}} V \sqrt{\frac{t}{n}} V \frac{t}{n} \right),$$

$$\tilde{\Lambda}_{k,\beta}(g; \Sigma; B) := \|B\|_1 \|Dg\|_{C^{k+1+\beta}} (\|\Sigma\| \vee \|\Sigma^{-1}\|) \log^2(2(\|\Sigma\| \vee \|\Sigma^{-1}\|)).$$

# Normal Approximation of $\langle \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle$

- Recall that

$$\begin{aligned} \langle \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle &= \langle D\mathcal{D}g_k(\Sigma)(\hat{\Sigma} - \Sigma), B \rangle \\ &+ \mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) + \langle \mathbb{E}\mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle \end{aligned}$$

- It follows from the concentration bound on the remainder that, for  $d = o(n)$ ,

$$|\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma) - \mathbb{E}\mathcal{S}_{\mathfrak{d}_k}(\Sigma; \hat{\Sigma} - \Sigma)| = o_{\mathbb{P}}(n^{-1/2})$$

# Normal Approximation of $\langle \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle$

- By the bound on the bias,

$$|\langle \mathbb{E} \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle| = O\left(\frac{d}{n}\right)^{(k+1+\beta)/2} = o(n^{-1/2})$$

provided that  $d \leq n^\alpha$ ,  $\alpha \in (0, 1)$  and  $s \geq k + 1 + \beta > \frac{1}{1-\alpha}$ .

- Thus

$$\langle \mathcal{D}g_k(\hat{\Sigma}) - \mathcal{D}g(\Sigma), B \rangle = \langle D\mathcal{D}g_k(\hat{\Sigma})(\hat{\Sigma} - \Sigma), B \rangle + o_{\mathbb{P}}(n^{-1/2}),$$

and asymptotic normality of  $n^{1/2} \langle D\mathcal{D}g_k(\hat{\Sigma})(\hat{\Sigma} - \Sigma), B \rangle$  follows from Berry-Esseen bound.

# Efficient Estimation of Linear Functionals of Principal Components (Koltchinskii, Löffler and Nickl (2017))

- $\lambda(\Sigma) = \|\Sigma\|$
- $\lambda(\Sigma)$  eigenvalue of multiplicity 1
- $g(\Sigma) := \text{dist}(\lambda(\Sigma); \sigma(\Sigma) \setminus \{\lambda(\Sigma)\})$
- $\theta(\Sigma)$  eigenvector of  $\Sigma$  corresponding to  $\lambda(\Sigma)$ ,  $\|\theta(\Sigma)\| = 1$  (the top principal component)
- **Problem:** given  $u \in \mathbb{H}$ , estimate  $\langle \theta(\Sigma), u \rangle$  based on i.i.d. observations  $X_1, \dots, X_n \sim N(0; \Sigma)$

# Efficient Estimation of Linear Functionals of Principal Components

- $\Sigma = \sum_{\lambda \in \sigma(\Sigma)} \lambda P_\lambda$
- $C(\Sigma) := \sum_{\lambda \in \sigma(\Sigma), \lambda \neq \lambda(\Sigma)} \frac{1}{\lambda - \lambda(\Sigma)} P_\lambda$
- $\sigma_u^2(\Sigma) := \|\Sigma\| \langle \Sigma C(\Sigma) u, C(\Sigma) u \rangle$
- For  $r > 1, a > 1, \sigma_0^2 > 0$ , denote

$$S(r; a) := \left\{ \Sigma : \mathbf{r}(\Sigma) \leq r, \frac{\|\Sigma\|}{g(\Sigma)} \leq a \right\}$$

# Efficient Estimation of Linear Functionals: Lower Bound

## Theorem

Let  $r > 1$ ,  $a > 1$ ,  $\sigma_0^2 > 0$ . Suppose, for some  $r' < r$ ,  $a' < a$  and  $\sigma'_0 > \sigma_0$

$$\mathcal{S}(r'; a') \cap \{\Sigma : \sigma_u(\Sigma) \geq \sigma'_0\} \neq \emptyset.$$

Then

$$\liminf_n \inf_{T_n} \sup_{\Sigma \in \mathcal{S}(r; a), \sigma_u(\Sigma) \geq \sigma_0} \frac{n \mathbb{E}_{\Sigma} \left( T_n(X_1, \dots, X_n) - \langle \theta(\Sigma), u \rangle \right)^2}{\sigma_u^2(\Sigma)} \geq 1,$$

where the infimum is taken over all sequences of estimators  $T_n = T_n(X_1, \dots, X_n)$ .

# Efficient Estimation of Linear Functionals: Asymptotic Normality

## Theorem

Suppose  $a > 1$ ,  $\sigma_0^2 > 0$  and  $r_n = o(n)$  as  $n \rightarrow \infty$ . There exists a sequence of estimators  $\hat{\theta}_n = \hat{\theta}_n(X_1, \dots, X_n) \in \mathbb{H}$  such that

$$\sup_{\Sigma \in \mathcal{S}(r_n; a), \sigma_u(\Sigma) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_\Sigma \left\{ \frac{n^{1/2} \left( \langle \hat{\theta}_n, u \rangle - \langle \theta(\Sigma), u \rangle \right)}{\sigma_u(\Sigma)} \leq x \right\} - \Phi(x) \right| \rightarrow 0$$

and, for all  $\ell \in \mathcal{L}$ ,

$$\sup_{\Sigma \in \mathcal{S}(r_n; a), \sigma_u(\Sigma) \geq \sigma_0} \left| \mathbb{E}_\Sigma \ell \left( \frac{n^{1/2} \left( \langle \hat{\theta}_n, u \rangle - \langle \theta(\Sigma), u \rangle \right)}{\sigma_u(\Sigma)} \right) - \mathbb{E} \ell(Z) \right| \rightarrow 0$$

as  $n \rightarrow \infty$ .



# Plug-in estimator $\theta(\hat{\Sigma}_n)$

## Theorem (Koltchinskii and Lounici (2016))

Suppose  $r_n = o(n)$ . Then

$$\sup_{\Sigma \in \mathcal{S}(r_n; a), \sigma_u(\Sigma) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_{\Sigma} \left\{ \zeta_n(\Sigma; u) \leq x \right\} - \Phi(x) \right| \rightarrow 0 \text{ as } n \rightarrow \infty,$$

where

$$\zeta_n(\Sigma; u) := \frac{n^{1/2} \left( \langle \theta(\hat{\Sigma}_n), u \rangle - d_n(\Sigma) \langle \theta(\Sigma), u \rangle \right)}{\sigma_u(\Sigma)}$$

and

$$d_n(\Sigma) := \mathbb{E}_{\Sigma}^{1/2} \langle \theta(\hat{\Sigma}_n), \theta(\Sigma) \rangle^2.$$

Moreover,  $d_n^2(\Sigma) = 1 + b_n(\Sigma)$  with  $|b_n(\Sigma)| \asymp \frac{\|\Sigma\|^2}{g^2(\Sigma)} \frac{r(\Sigma)}{n}$ .

# Plug-in estimator $\theta(\hat{\Sigma}_n)$

- $\langle \theta(\hat{\Sigma}_n), u \rangle$  “concentrates around”  $d_n(\Sigma) \langle \theta(\Sigma), u \rangle$
- Its “bias”  $(d_n(\Sigma) - 1) \langle \theta(\Sigma), u \rangle$  could be as large as  $\frac{r_n}{n}$
- If  $\sqrt{n} \leq r_n = o(n)$ , the bias is too large.

# Bias Reduction (Koltchinskii and Lounici (2016))

- Suppose  $n = 2n'$
- $\hat{\Sigma}^{(1)}, \hat{\Sigma}^{(2)}$  are sample covariances based on two sub-samples of size  $n'$
- $\hat{b}_n := \langle \theta(\hat{\Sigma}^{(1)}), \theta(\hat{\Sigma}^{(2)}) \rangle - 1$
- For some  $\delta_n \rightarrow 0$ ,  $\sup_{\Sigma \in \mathcal{S}(r_n; a)} \mathbb{P}_{\Sigma} \left\{ |\hat{b}_n - b_{n/2}(\Sigma)| \geq \frac{\delta_n}{\sqrt{n}} \right\} \rightarrow 0$
- Let  $\hat{d}_n := \sqrt{1 + \hat{b}_n}$  and  $\tilde{\theta}_n := \theta(\hat{\Sigma}^{(1)})/\hat{d}_n$ . Then, under the assumption  $r_n = o(n)$ ,

$$\sup_{\Sigma \in \mathcal{S}(r_n; a), \sigma_u(\Sigma) \geq \sigma_0} \sup_{x \in \mathbb{R}} \left| \mathbb{P}_{\Sigma} \left\{ \frac{\sqrt{\frac{n}{2}} \left( \langle \tilde{\theta}_n, u \rangle - \langle \theta(\Sigma), u \rangle \right)}{\sigma_u(\Sigma)} \leq x \right\} - \Phi(x) \right| \rightarrow 0.$$

# Construction of Asymptotically Efficient Estimators $\hat{\theta}_n$ .

- Recall that  $r_n = o(n)$
- Let  $m = m_n$  be such that  $m_n = o(n)$  and  $r_n = o(m_n)$ .
- Split the sample  $X_1, \dots, X_n$  into three sub-samples, the first one of size  $n' = n - 2m$  and two others of size  $m$  each and construct three sample covariances,  $\hat{\Sigma}^{(1)}$ ,  $\hat{\Sigma}^{(2)}$  and  $\hat{\Sigma}^{(3)}$ , based on each of the sub-samples.
- Finally, define

$$\tilde{d}_n := \frac{\langle \theta(\hat{\Sigma}^{(1)}), \theta(\hat{\Sigma}^{(2)}) \rangle}{\langle \theta(\hat{\Sigma}^{(2)}), \theta(\hat{\Sigma}^{(3)}) \rangle^{1/2}}$$

and

$$\hat{\theta}_n := \frac{\theta(\hat{\Sigma}^{(1)})}{\tilde{d}_n}.$$